

On \mathcal{Z}_p -norms of random vectors

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Abstract

To any n -dimensional random vector X we may associate its L_p -centroid body $\mathcal{Z}_p(X)$ and the corresponding norm. We formulate a conjecture concerning the bound on the $\mathcal{Z}_p(X)$ -norm of X and show that it holds under some additional symmetry assumptions. We also relate our conjecture with estimates of covering numbers and Sudakov-type minorization bounds.

1 Introduction. Formulation of the Problem.

Let $p \geq 2$ and $X = (X_1, \dots, X_n)$ be a random vector in \mathbb{R}^n such that $\mathbb{E}|X|^p < \infty$. We define the following two norms on \mathbb{R}^n :

$$\|t\|_{\mathcal{M}_p(X)} := (\mathbb{E}|\langle t, X \rangle|^p)^{1/p} \quad \text{and} \quad \|t\|_{\mathcal{Z}_p(X)} := \sup\{|\langle t, s \rangle| : \|s\|_{\mathcal{M}_p(X)} \leq 1\}.$$

By $\mathcal{M}_p(X)$ and $\mathcal{Z}_p(X)$ we will also denote unit balls in these norms, i.e.

$$\mathcal{M}_p(X) := \{t \in \mathbb{R}^n : \|t\|_{\mathcal{M}_p(X)} \leq 1\} \quad \text{and} \quad \mathcal{Z}_p(X) := \{t \in \mathbb{R}^n : \|t\|_{\mathcal{Z}_p(X)} \leq 1\}.$$

The set $\mathcal{Z}_p(X)$ is called the L_p -centroid body of X (or rather of the distribution of X). It was introduced (under a different normalization) for uniform distributions on convex bodies in [9]. Investigation of L_p -centroid bodies played a crucial role in the Paouris proof of large deviations bounds for Euclidean norms of log-concave vectors [10]. Such bodies also appears in questions related to the optimal concentration of log-concave vectors [7].

Let us introduce a bit of useful notation. We set $|t| := \|t\|_2 = \sqrt{\langle t, t \rangle}$ and $B_2^n = \{t \in \mathbb{R}^n : |t| \leq 1\}$. By $\|Y\|_p = (\mathbb{E}|Y|^p)^{1/p}$ we denote the L_p -norm of a random variable Y . Letter C denotes universal constants (that may differ at each occurrence), we write $f \sim g$ if $\frac{1}{C}f \leq g \leq Cf$.

Let us begin with a simple case, when a random vector X is rotationally invariant. Then $X = RU$, where U has a uniform distribution on S^{n-1} and $R = |X|$ is a nonnegative random variable, independent of U . We have for any vector $t \in \mathbb{R}^n$ and $p \geq 2$,

$$\|\langle t, U \rangle\|_p = |t| \|U_1\|_p \sim \sqrt{\frac{p}{n+p}} |t|,$$

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where U_1 is the first coordinate of U . Therefore

$$\|t\|_{\mathcal{M}_p(X)} = \|U_1\|_p \|R\|_p |t| \quad \text{and} \quad \|t\|_{\mathcal{Z}_p(X)} = \|U_1\|_p^{-1} \|R\|_p^{-1} |t|.$$

So

$$\left(\mathbb{E}\|X\|_{\mathcal{Z}_p(X)}^p\right)^{1/p} = \|U_1\|_p^{-1} \|R\|_p^{-1} (\mathbb{E}|X|^p)^{1/p} = \|U_1\|_p^{-1} \sim \sqrt{\frac{n+p}{p}}. \quad (1)$$

This motivates the following problem.

Problem 1. Is it true that for (at least a large class of) centered n -dimensional random vectors X ,

$$\left(\mathbb{E}\|X\|_{\mathcal{Z}_p(X)}^2\right)^{1/2} \leq C \sqrt{\frac{n+p}{p}} \quad \text{for } p \geq 2,$$

or maybe even

$$\left(\mathbb{E}\|X\|_{\mathcal{Z}_p(X)}^p\right)^{1/p} \leq C \sqrt{\frac{n+p}{p}} \quad \text{for } p \geq 2?$$

Notice that the problem is linearly-invariant, since

$$\|AX\|_{\mathcal{Z}_p(AX)} = \|X\|_{\mathcal{Z}_p(X)} \quad \text{for any } A \in \text{GL}(n). \quad (2)$$

For any centered random vector X with nondegenerate covariance matrix, random vector $Y = \text{Cov}(X)^{-1/2}X$ is isotropic (i.e. centered with identity covariance matrix). We have $\mathcal{M}_2(Y) = \mathcal{Z}_2(Y) = B_2^n$, hence

$$\mathbb{E}\|X\|_{\mathcal{Z}_2(X)}^2 = \mathbb{E}\|Y\|_{\mathcal{Z}_2(Y)}^2 = \mathbb{E}|Y|^2 = n.$$

Next remark shows that the answer to our problem is positive in the case $p \geq n$.

Remark 1. For $p \geq n$ and any n -dimensional random vector X we have $(\mathbb{E}\|X\|_{\mathcal{Z}_p(X)}^p)^{1/p} \leq 10$.

Proof. Let S be a $1/2$ -net in the unit ball of $\mathcal{M}_p(X)$ such that $|S| \leq 5^n$ (such net exists by the volume-based argument, cf. [1, Corollary 4.1.15]). Then

$$\begin{aligned} (\mathbb{E}\|X\|_{\mathcal{Z}_p(X)}^p)^{1/p} &\leq 2 \left(\mathbb{E} \sup_{t \in S} |\langle t, X \rangle|^p \right)^{1/p} \leq 2 \left(\mathbb{E} \sum_{t \in S} |\langle t, X \rangle|^p \right)^{1/p} \\ &\leq 2|S|^{1/p} \sup_{t \in S} (\mathbb{E} |\langle t, X \rangle|^p)^{1/p} \leq 2 \cdot 5^{n/p}. \end{aligned}$$

□

L_p -centroid bodies play an important role in the study of vectors uniformly distributed on convex bodies and a more general class of log-concave vectors. A random vector with a nondegenerate covariance matrix is called log-concave if its density has the form e^{-h} , where $h: \mathbb{R}^n \rightarrow (-\infty, \infty]$ is convex. If X is centered and log-concave then

$$\|\langle t, X \rangle\|_p \leq \lambda \frac{p}{q} \|\langle t, X \rangle\|_q \quad \text{for } p \geq q \geq 2, \quad (3)$$

where $\lambda = 2$ ($\lambda = 1$ if X is symmetric and log-concave and $\lambda = 3$ for arbitrary log-concave vectors). One of open problems for log-concave vectors [7] states that for such vectors, arbitrary norm $\|\cdot\|$ and $q \geq 1$,

$$(\mathbb{E}\|X\|^q)^{1/q} \leq C \left(\mathbb{E}\|X\| + \sup_{\|t\|_* \leq 1} \|\langle t, X \rangle\|_q \right).$$

In particular one may expect that for log-concave vectors

$$(\mathbb{E}\|X\|_{\mathcal{Z}_p(X)}^q)^{1/q} \leq C \left(\mathbb{E}\|X\|_{\mathcal{Z}_p(X)} + \sup_{t \in \mathcal{M}_p(X)} \|\langle t, X \rangle\|_q \right) \leq C \left(\mathbb{E}\|X\|_{\mathcal{Z}_p(X)} + \frac{\max\{p, q\}}{p} \right).$$

As a result it is natural to state the following variant of Problem 1.

Problem 2. Let X be a centered log-concave n -dimensional random vector. Is it true that

$$(\mathbb{E}\|X\|_{\mathcal{Z}_p(X)}^q)^{1/q} \leq C \sqrt{\frac{n}{p}} \quad \text{for } 2 \leq p \leq n, \quad 1 \leq q \leq \sqrt{pn}.$$

In Section 2 we show that Problems 1 and 2 have affirmative solutions in the class of unconditional vectors. In Section 3 we relate our problems to estimates of covering numbers. We also show that the first estimate in Problem 1 holds if the random vector X satisfies the Sudakov-type minorization bound.

2 Bounds for unconditional random vectors

In this section we consider the class of *unconditional* random vectors in \mathbb{R}^n , i.e. vectors X having the same distribution as $(\varepsilon_1|X_1|, \varepsilon_2|X_2|, \dots, \varepsilon_n|X_n|)$, where (ε_i) is a sequence of independent symmetric ± 1 random variables (Rademacher sequence), independent of X .

Our first result shows that formula (1) may be extended to the unconditional case for p even. We use the standard notation – for a multiindex $\alpha = (\alpha_1, \dots, \alpha_n)$, $x \in \mathbb{R}^n$ and $m = \sum \alpha_i$, $x^\alpha := \prod_i x_i^{\alpha_i}$ and $\binom{m}{\alpha} := m! / (\prod_i \alpha_i!)$.

Proposition 2. *We have for any $k = 1, 2, \dots$ and any n -dimensional unconditional random vector X such that $\mathbb{E}|X|^{2k} < \infty$,*

$$\left(\mathbb{E}\|X\|_{\mathcal{Z}_{2k}(X)}^{2k}\right)^{1/(2k)} \leq c_{2k} := \left(\sum_{\|\alpha\|_1=k} \frac{\binom{k}{\alpha}^2}{\binom{2k}{2\alpha}}\right)^{1/(2k)} \sim \sqrt{\frac{n+k}{k}},$$

where the summation runs over all multiindices $\alpha = (\alpha_1, \dots, \alpha_n)$ with nonnegative integer coefficients such that $\|\alpha\|_1 = \sum_{i=1}^n \alpha_i = k$.

Proof. Observe first that

$$\mathbb{E}|\langle t, X \rangle|^{2k} = \mathbb{E}\left|\sum_{i=1}^n t_i \varepsilon_i X_i\right|^{2k} = \sum_{\|\alpha\|_1=k} \binom{2k}{2\alpha} t^{2\alpha} \mathbb{E}X^{2\alpha}.$$

For any $t, s \in \mathbb{R}^n$ we have

$$|\langle t, s \rangle|^k = \sum_{\|\alpha\|_1=k} \binom{k}{\alpha} t^\alpha s^\alpha.$$

So by the Cauchy-Schwarz inequality,

$$\|s\|_{\mathcal{Z}_{2k}(X)}^k = \sup\{|\langle t, s \rangle|^k : \mathbb{E}|\langle t, X \rangle|^{2k} \leq 1\} \leq \left(\sum_{\|\alpha\|_1=k} \frac{\binom{k}{\alpha}^2}{\binom{2k}{2\alpha}} \frac{s^{2\alpha}}{\mathbb{E}X^{2\alpha}}\right)^{1/2}.$$

To see that $c_{2k} \sim \sqrt{(n+k)/k}$ observe that

$$\frac{\binom{k}{\alpha}^2}{\binom{2k}{2\alpha}} = \binom{2k}{k}^{-1} \prod_{i=1}^n \binom{2\alpha_i}{\alpha_i}.$$

Therefore, since $1 \leq \binom{2l}{l} \leq 2^{2l}$, we get

$$4^{-k} \binom{n+k-1}{k} \leq c_{2k}^2 \leq 4^k \binom{n+k-1}{k}.$$

□

Corollary 3. *Let X be an unconditional n -dimensional random vector. Then*

$$\left(\mathbb{E}\|X\|_{\mathcal{Z}_p(X)}^{2k}\right)^{1/2k} \leq C \sqrt{\frac{n+p}{p}} \quad \text{for any positive integer } k \leq \frac{p}{2}.$$

Proof. By the monotonicity of L_{2k} -norms we may and will assume that $k = \lfloor p/2 \rfloor$. Then by Proposition 2,

$$\left(\mathbb{E}\|X\|_{\mathcal{Z}_p(X)}^{2k}\right)^{1/2k} \leq \left(\mathbb{E}\|X\|_{\mathcal{Z}_{2k}(X)}^{2k}\right)^{1/2k} \leq C\sqrt{\frac{n+k}{k}} \leq C\sqrt{\frac{n+p}{p}}.$$

□

In the unconditional log-concave case we may bound higher moments of $\|X\|_{\mathcal{Z}_p(X)}$.

Theorem 4. *Let X be an unconditional log-concave n -dimensional random vector. Then for $p, q \geq 2$,*

$$\left(\mathbb{E}\|X\|_{\mathcal{Z}_p(X)}^q\right)^{1/q} \leq C\left(\sqrt{\frac{n+p}{p}} + \sup_{t \in \mathcal{M}_p(X)} \|\langle t, X \rangle\|_q\right) \leq C\left(\sqrt{\frac{n+p}{p}} + \frac{q}{p}\right).$$

In order to show this result we will need the following lemma.

Lemma 5. *Let $2 \leq p \leq n$, X be an unconditional random vector in \mathbb{R}^n such that $\mathbb{E}|X|^p < \infty$ and $\mathbb{E}|X_i| = 1$. Then*

$$\|s\|_{\mathcal{Z}_p(X)} \leq \sup_{I \subset [n], |I| \leq p} \sup_{\|t\|_{\mathcal{M}_p(X)} \leq 1} \left| \sum_{i \in I} t_i s_i \right| + C_1 \sup_{\|t\|_{\mathcal{M}_p(X)} \leq 1, \|t\|_2 \leq p^{-1/2}} \left| \sum_{i=1}^n t_i s_i \right|. \quad (4)$$

Proof. We have by the unconditionality of X and Jensen's inequality,

$$\|t\|_{\mathcal{M}_p(X)} = \left\| \sum_{i=1}^n t_i \varepsilon_i |X_i| \right\|_p \geq \left\| \sum_{i=1}^n t_i \varepsilon_i \mathbb{E}|X_i| \right\|_p.$$

By the result of Hitczenko [5], for numbers a_1, \dots, a_n ,

$$\left\| \sum_{i=1}^n a_i \varepsilon_i \right\|_p \sim \sum_{i \leq p} a_i^* + \sqrt{p} \left(\sum_{i > p} |a_i^*|^2 \right)^{1/2}, \quad (5)$$

where $(a_i^*)_{i \leq n}$ denotes the nonincreasing rearrangement of $(|a_i|)_{i \leq n}$. Thus

$$\sqrt{p} \left(\sum_{i > p} |t_i^*|^2 \right)^{1/2} \leq C_1 \|t\|_{\mathcal{M}_p(X)}$$

and (4) easily follows. □

Proof of Theorem 4. The last bound in the assertion follows by (3). It is easy to see that (increasing q if necessary) it is enough to consider the case $q \geq \sqrt{np}$.

If $q \geq n$ then the similar argument as in the proof of Remark 1 shows that

$$\left(\mathbb{E}\|X\|_{\mathcal{Z}_p(X)}^q\right)^{1/q} \leq 2 \cdot 5^{n/q} \sup_{t \in \mathcal{M}_p(X)} \|\langle t, X \rangle\|_q \leq 10 \sup_{t \in \mathcal{M}_p(X)} \|\langle t, X \rangle\|_q.$$

Finally, consider the remaining case $\sqrt{pn} \leq q \leq n$. By (2) we may assume that $\mathbb{E}|X_i| = 1$ for all i . By the log-concavity $\|\langle t, X \rangle\|_{q_1} \leq C \frac{q_1}{q_2} \|\langle t, X \rangle\|_{q_2}$ for $q_1 \geq q_2 \geq 1$, in particular $\sigma_i := \|X_i\|_2 \leq C$.

Let $\mathcal{E}_1, \dots, \mathcal{E}_n$ be i.i.d. symmetric exponential random variables with variance 1. By [6, Theorem 3.1] we have

$$\begin{aligned} & \left\| \sup_{\|t\|_{\mathcal{M}_p(X)} \leq 1, \|t\|_2 \leq p^{-1/2}} \left| \sum_{i=1}^n t_i X_i \right| \right\|_q \\ & \leq C \left(\left\| \sup_{\|t\|_{\mathcal{M}_p(X)} \leq 1, \|t\|_2 \leq p^{-1/2}} \left| \sum_{i=1}^n t_i \sigma_i \mathcal{E}_i \right| \right\|_1 + \sup_{\|t\|_{\mathcal{M}_p(X)} \leq 1, \|t\|_2 \leq p^{-1/2}} \|\langle t, X \rangle\|_q \right). \end{aligned}$$

We have

$$\sup_{\|t\|_{\mathcal{M}_p(X)} \leq 1, \|t\|_2 \leq p^{-1/2}} \|\langle t, X \rangle\|_q \leq \sup_{\|t\|_{\mathcal{M}_p(X)} \leq 1} \|\langle t, X \rangle\|_q$$

and

$$\left\| \sup_{\|t\|_{\mathcal{M}_p(X)} \leq 1, \|t\|_2 \leq p^{-1/2}} \left| \sum_{i=1}^n t_i \sigma_i \mathcal{E}_i \right| \right\|_1 \leq \frac{1}{\sqrt{p}} \left\| \sqrt{\sum_{i=1}^n \sigma_i^2 \mathcal{E}_i^2} \right\|_1 \leq \frac{1}{\sqrt{p}} \sqrt{\sum_{i=1}^n \sigma_i^2} \leq C \sqrt{\frac{n}{p}}.$$

Thus

$$\left\| \sup_{\|t\|_{\mathcal{M}_p(X)} \leq 1, \|t\|_2 \leq p^{-1/2}} \left| \sum_{i=1}^n t_i X_i \right| \right\|_q \leq C \left(\sqrt{\frac{n}{p}} + \sup_{\|t\|_{\mathcal{M}_p(X)} \leq 1} \|\langle t, X \rangle\|_q \right).$$

Let for each $I \subset [n]$, $P_I X = (X_i)_{i \in I}$ and S_I be a 1/2-net in $\mathcal{M}_p(P_I X)$ of cardinality at

most $5^{|I|}$. We have

$$\begin{aligned}
\left\| \sup_{I \subset [n], |I| \leq p} \sup_{\|t\|_{\mathcal{M}_p(X)} \leq 1} \left| \sum_{i \in I} t_i X_i \right| \right\|_q &\leq 2 \left\| \sup_{I \subset [n], |I| \leq p} \sup_{t \in S_I} \left| \sum_{i \in I} t_i X_i \right| \right\|_q \\
&\leq 2 \left(\sum_{I \subset [n], |I| \leq p} \sum_{t \in S_I} \mathbb{E} \left| \sum_{i \in I} t_i X_i \right|^q \right)^{1/q} \\
&\leq 2 \cdot 5^{p/q} |\{I \subset [n], |I| \leq p\}|^{1/q} \sup_I \sup_{t \in S_I} \left\| \sum_{i \in I} t_i X_i \right\|_q \\
&\leq 10 \left(\frac{en}{p} \right)^{p/q} \sup_{t \in \mathcal{M}_p(X)} \left\| \sum_{i \in I} t_i X_i \right\|_q \\
&\leq C \sup_{t \in \mathcal{M}_p(X)} \left\| \sum_{i \in I} t_i X_i \right\|_q,
\end{aligned}$$

where the last estimate follows from $q \geq \sqrt{np}$.

Hence the assertion follows by Lemma 5. \square

Corollary 6. *Let X be an unconditional log-concave n -dimensional random vector and $2 \leq p \leq n$. Then*

$$\frac{1}{C} \sqrt{\frac{n}{p}} \leq \mathbb{E} \|X\|_{\mathcal{Z}_p(X)} \leq \left(\mathbb{E} \|X\|_{\mathcal{Z}_p(X)}^{\sqrt{np}} \right)^{1/\sqrt{np}} \leq C \sqrt{\frac{n}{p}} \quad (6)$$

and

$$\mathbb{P} \left(\|X\|_{\mathcal{Z}_p(X)} \geq \frac{1}{C} \sqrt{\frac{n}{p}} \right) \geq \frac{1}{C}, \quad \mathbb{P} \left(\|X\|_{\mathcal{Z}_p(X)} \geq Ct \sqrt{\frac{n}{p}} \right) \leq e^{-t\sqrt{np}} \text{ for } t \geq 1.$$

Proof. The upper bound in (6) easily follows by Theorem 4. In fact we have for $t \geq 1$,

$$\left(\mathbb{E} \|X\|_{\mathcal{Z}_p(X)}^{t\sqrt{np}} \right)^{1/(t\sqrt{np})} \leq Ct \sqrt{\frac{n}{p}},$$

hence the Chebyshev inequality yields the upper tail bound for $\|X\|_{\mathcal{Z}_p(X)}$.

To establish lower bounds we may assume that X is additionally isotropic. Then by the result of Bobkov and Nazarov [3] we have $\| \langle t, X \rangle \|_p \leq C(\sqrt{p} \|t\|_2 + p \|t\|_\infty)$. This easily gives

$$\mathbb{E} \|X\|_{\mathcal{Z}_p(X)} \geq \frac{1}{C} \sqrt{\frac{n}{p}} \mathbb{E} X_{\lceil n/2 \rceil}^* \geq \frac{1}{C} \sqrt{\frac{n}{p}},$$

where the last inequality follows by Lemma 7 below.

By the Paley-Zygmund inequality we get

$$\mathbb{P}\left(\|X\|_{\mathcal{Z}_p(X)} \geq \frac{1}{C}\sqrt{\frac{n}{p}}\right) \geq \mathbb{P}\left(\|X\|_{\mathcal{Z}_p(X)} \geq \frac{1}{2}\mathbb{E}\|X\|_{\mathcal{Z}_p(X)}\right) \geq \frac{(\mathbb{E}\|X\|_{\mathcal{Z}_p(X)})^2}{4\mathbb{E}\|X\|_{\mathcal{Z}_p(X)}^2} \geq c.$$

□

Lemma 7. *Let X be a symmetric isotropic n -dimensional log-concave vector. Then $\mathbb{E}X_{[n/2]}^* \geq \frac{1}{C}$.*

Proof. Let $a_i > 0$ be such that $\mathbb{P}(X_i \geq a_i) = 3/8$. Then by the log-concavity of X_i , $\mathbb{P}(|X_i| \geq ta_i) = 2\mathbb{P}(X_i \geq ta_i) \leq (3/4)^t$ for $t \geq 1$ and integration by parts yields $\|X_i\|_2 \leq Ca_i$. Thus $a_i \geq c_1$ for a constant $c_1 > 0$.

Let $S = \sum_{i=1}^n I_{\{|X_i| \geq c_1\}}$. Then $\mathbb{E}S = \sum_{i=1}^n \mathbb{P}(|X_i| \geq c_1) \geq 3n/4$. On the other hand $\mathbb{E}S \leq \frac{n}{2} + n\mathbb{P}(X_{[n/2]}^* \geq c_1)$, so

$$\mathbb{E}X_{[n/2]}^* \geq c_1\mathbb{P}(X_{[n/2]}^* \geq c_1) \geq c_1/4.$$

□

The next example shows that the tail and moment bounds in Corollary 6 are optimal.

Example. Let $X = (X_1, \dots, X_n)$ be an isotropic random vector with i.i.d. symmetric exponential coordinates (i.e. X has the density $2^{n/2} \exp(-\sqrt{2}\|x\|_1)$). Then $(\mathbb{E}|X_i|^p)^{1/p} \leq p/2$, so $\frac{2}{p}e_i \in \mathcal{M}_p(X)$ and

$$\mathbb{P}\left(\|X\|_{\mathcal{Z}_p(X)} \geq t\sqrt{n/p}\right) \geq \mathbb{P}(|X_i| \geq t\sqrt{np}/2) \geq e^{-t\sqrt{np}/\sqrt{2}}$$

and for $q = s\sqrt{np}$, $s \geq 1$,

$$\left(\mathbb{E}\|X\|_{\mathcal{Z}_p(X)}^q\right)^{1/q} \geq \frac{2}{p}\|X_i\|_q \geq cq/p = cs\sqrt{n/p}.$$

3 General case – approach via entropy numbers

In this section we propose a method of deriving estimates for \mathcal{Z}_p -norms via entropy estimates for \mathcal{M}_p -balls and Euclidean distance. We use a standard notation – for sets $T, S \subset \mathbb{R}^n$, by $N(T, S)$ we denote the minimal number of translates of S that are enough to cover T . If S is the ε -ball with respect to some translation-invariant metric d then $N(T, S)$ is also denoted as $N(T, d, \varepsilon)$ and is called the metric entropy of T with respect to d .

We are mainly interested in log-concave vectors or random vectors which satisfy moment estimates

$$\|\langle t, X \rangle\|_p \leq \lambda \frac{p}{q} \|\langle t, X \rangle\|_q \quad \text{for } p \geq q \geq 2. \quad (7)$$

Let us start with a simple bound.

Proposition 8. *Suppose that X is isotropic in \mathbb{R}^n and (7) holds. Then for any $p \geq 2$ and $\varepsilon > 0$ we have*

$$\left(\mathbb{E}\|X\|_{\mathcal{Z}_p(X)}^2\right)^{1/2} \leq \varepsilon\sqrt{n} + \frac{e\lambda}{p} \max\{p, \log N(\mathcal{M}_p(X), \varepsilon B_2^n)\}.$$

Proof. Let $N = N(\mathcal{M}_p(X), \varepsilon B_2^n)$ and choose $t_1, \dots, t_N \in \mathcal{M}_p(X)$ such that $\mathcal{M}_p(X) \subset \bigcup_{i=1}^N (t_i + \varepsilon B_2^n)$. Then

$$\|x\|_{\mathcal{Z}_p(X)} \leq \varepsilon|x| + \sup_{i \leq N} \langle t_i, x \rangle.$$

Let $r = \max\{p, \log N\}$. We have

$$\begin{aligned} \left(\mathbb{E} \sup_{i \leq N} |\langle t_i, X \rangle|^2\right)^{1/2} &\leq \left(\mathbb{E} \sup_{i \leq N} |\langle t_i, X \rangle|^r\right)^{1/r} \leq \left(\sum_{i=1}^N \mathbb{E} |\langle t_i, X \rangle|^r\right)^{1/r} \\ &\leq N^{1/r} \sup_i \|\langle t_i, X \rangle\|_r \leq e\lambda \frac{r}{p} \sup_i \|\langle t_i, X \rangle\|_p \leq e\lambda \frac{r}{p}. \end{aligned}$$

□

Remark 9. The Paouris inequality [10] states that for isotropic log-concave vectors and $q \geq 2$, $(\mathbb{E}|X|^q)^{1/q} \leq C(\sqrt{n} + q)$, so for such vectors and $q \geq 2$,

$$\left(\mathbb{E}\|X\|_{\mathcal{Z}_p(X)}^q\right)^{1/q} \leq C\varepsilon(\sqrt{n} + q) + \frac{2e}{p} \max\{p, q, \log N(\mathcal{M}_p(X), \varepsilon B_2^n)\}.$$

Unfortunately, the known estimates for entropy numbers of \mathcal{M}_p -balls are rather weak.

Theorem 10 ([4, Proposition 9.2.8]). *Assume that X is isotropic log-concave and $2 \leq p \leq \sqrt{n}$. Then*

$$\log N\left(\mathcal{M}_p(X), \frac{t}{\sqrt{p}} B_2^n\right) \leq C \frac{n \log^2 p \log t}{t} \quad \text{for } 1 \leq t \leq \min\left\{\sqrt{p}, \frac{1}{C} \frac{n \log p}{p^2}\right\}.$$

Corollary 11. *Let X be isotropic log-concave, then*

$$\left(\mathbb{E}\|X\|_{\mathcal{Z}_p(X)}^p\right)^{1/p} \leq C \left(\frac{n}{p}\right)^{3/4} \log p \sqrt{\log n} \quad \text{for } 2 \leq p \leq \frac{1}{C} n^{3/7} \log^{-2/7} n.$$

Proof. We apply Theorem 10 with $t = (n/p)^{1/4} \log p \log^{1/2} n$ and Proposition 8 with $\varepsilon = tp^{-1/2}$. □

Remark 12. Suppose that X is centered and the following stronger bound than (7) (satisfied for example for Gaussian vectors) holds

$$\|\langle t, X \rangle\|_p \leq \lambda \sqrt{\frac{p}{q}} \|\langle t, X \rangle\|_q \quad \text{for } p \geq q \geq 2. \quad (8)$$

Then for any $2 \leq p \leq n$,

$$\frac{1}{\lambda} \sqrt{\frac{2n}{p}} \leq \left(\mathbb{E} \|X\|_{\mathcal{Z}_p(X)}^2 \right)^{1/2} \leq \left(\mathbb{E} \|X\|_{\mathcal{Z}_p(X)}^n \right)^{1/n} \leq 10\lambda \sqrt{\frac{n}{p}}.$$

Proof. Without loss of generality we may assume that X is isotropic. We have

$$\|\langle t, X \rangle\|_p \leq \lambda \sqrt{p/2} \|\langle t, X \rangle\|_2 = \lambda \sqrt{p/2} |t|,$$

so $\mathcal{M}_p(X) \supset \lambda^{-1} \sqrt{2/p} B_2^n$ and

$$\left(\mathbb{E} \|X\|_{\mathcal{Z}_p(X)}^2 \right)^{1/2} \geq \frac{1}{\lambda} \sqrt{\frac{2}{p}} (\mathbb{E} |X|^2)^{1/2} = \frac{1}{\lambda} \sqrt{\frac{2n}{p}}.$$

On the other hand let S be a $1/2$ -net in $\mathcal{M}_p(X)$ of cardinality at most 5^n . Then

$$\begin{aligned} \left(\mathbb{E} \|X\|_{\mathcal{Z}_p(X)}^n \right)^{1/n} &\leq 2 \left(\mathbb{E} \sup_{t \in S} |\langle t, X \rangle|^n \right)^{1/n} \leq 2 \left(\sum_{t \in S} \mathbb{E} |\langle t, X \rangle|^n \right)^{1/n} \\ &\leq 2|S|^{1/n} \sup_{t \in S} \|\langle t, X \rangle\|_n \leq 10\lambda \sqrt{\frac{n}{p}} \sup_{t \in S} \|\langle t, X \rangle\|_p \leq 10\lambda \sqrt{\frac{n}{p}}. \end{aligned}$$

□

Recall that the Sudakov minoration principle [11] states that if G is an isotropic Gaussian vector in \mathbb{R}^n then for any bounded $T \subset \mathbb{R}^n$ and $\varepsilon > 0$,

$$\mathbb{E} \sup_{t \in T} \langle t, G \rangle \geq \frac{1}{C} \varepsilon \sqrt{\log N(T, \varepsilon B_2^n)}.$$

So we can say that a random vector X in \mathbb{R}^n satisfies *the L_2 -Sudakov minoration with a constant C_X* if for any bounded $T \subset \mathbb{R}^n$ and $\varepsilon > 0$,

$$\mathbb{E} \sup_{t \in T} \langle t, X \rangle \geq \frac{1}{C_X} \varepsilon \sqrt{\log N(T, \varepsilon B_2^n)}.$$

Example. Any unconditional n -dimensional random vector satisfies the L_2 -Sudakov minoration with constant $C \sqrt{\log(n+1)} / (\min_{i \leq n} \mathbb{E} |X_i|)$.

Indeed, we have by the unconditionality, Jensen's inequality and the contraction principle,

$$\mathbb{E} \sup_{t \in T} \sum_{i=1}^n t_i X_i = \mathbb{E} \sup_{t \in T} \sum_{i=1}^n t_i \varepsilon_i |X_i| \geq \mathbb{E} \sup_{t \in T} \sum_{i=1}^n t_i \varepsilon_i \mathbb{E} |X_i| \geq \min_{i \leq n} \mathbb{E} |X_i| \mathbb{E} \sup_{t \in T} \sum_{i=1}^n t_i \varepsilon_i.$$

On the other hand, the classical Sudakov minoration and the contraction principle yields

$$\begin{aligned} \frac{1}{C} \varepsilon \sqrt{\log N(T, \varepsilon B_2^n)} &\leq \mathbb{E} \sup_{t \in T} \sum_{i=1}^n t_i g_i \leq \mathbb{E} \max_{i \leq n} |g_i| \mathbb{E} \sup_{t \in T} \sum_{i=1}^n t_i \varepsilon_i \\ &\leq C \sqrt{\log(n+1)} \mathbb{E} \sup_{t \in T} \sum_{i=1}^n t_i \varepsilon_i. \end{aligned}$$

However the L_2 -Sudakov minoration constant may be quite large in the isotropic case even for unconditional vectors if we do not assume that L_1 and L_2 norms of X_i are comparable. Indeed, let $\mathbb{P}(X = \pm n^{1/2} e_i) = \frac{1}{2n}$ for $i = 1, \dots, n$, where e_1, \dots, e_n is the canonical basis of \mathbb{R}^n . Then X is isotropic and unconditional. Let $T = \{t \in \mathbb{R}^n : \|t\|_\infty \leq n^{-1/2}\}$. Then

$$\mathbb{E} \sup_{t \in T} |\langle t, X \rangle| \leq 1.$$

However, by the volume-based estimate,

$$N(T, \varepsilon B_2^n) \geq \frac{\text{vol}(T)}{\text{vol}(\varepsilon B_2^n)} \geq \left(\frac{1}{C\varepsilon}\right)^n,$$

hence

$$\sup_{\varepsilon > 0} \varepsilon \sqrt{\log N((T, \varepsilon B_2^n))} \geq \frac{1}{C} \sqrt{n}.$$

Thus the L_2 -Sudakov constant $C_X \geq \sqrt{n}/C$ in this case.

Next proposition shows that random vectors with uniformly log-convex density satisfy the L_2 -Sudakov minoration.

Proposition 13. *Suppose that a symmetric random vector X in \mathbb{R}^n has the density of the form e^h such that $\text{Hess}(h) \geq -\alpha \text{Id}$ for some $\alpha > 0$. Then X satisfies the L_2 -Sudakov minoration with constant $C_X \leq C\sqrt{\alpha}$.*

Proof. We will follow the method of the proof of the (dual) classical Sudakov inequality (cf. (3.15) and its proof in [8]).

Let T be a bounded symmetric set and

$$A := \mathbb{E} \sup_{t \in T} |\langle t, X \rangle|.$$

By the duality of entropy numbers [2] we need to show that $\log^{1/2} N(\varepsilon^{-1} B_2^n, T^\circ) \leq C\varepsilon^{-1} \alpha^{1/2} A$ for $\varepsilon > 0$ or equivalently that

$$N(\delta B_2^n, 6AT^\circ) \leq \exp(C\alpha\delta^2) \quad \text{for } \delta > 0. \quad (9)$$

To this end let $N = N(\delta B_2^n, 6AT^0)$. If $N = 1$ there is nothing to show, so assume that $N \geq 2$. Then we may choose $t_1, \dots, t_N \in \delta B_2^n$ such that the balls $t_i + 3AT^0$ are disjoint. Let $\mu = \mu_X$ be the distribution of X . By the Chebyshev inequality,

$$\mu(3AT^0) = 1 - \mathbb{P}\left(\sup_{t \in T} |\langle t, X \rangle| > 3A\right) \geq \frac{2}{3}.$$

Observe also that for any symmetric set K and $t \in \mathbb{R}^n$,

$$\begin{aligned} \mu(t + K) &= \int_K e^{h(x-t)} dx = \int_K e^{h(x+t)} dx = \int_K \frac{1}{2}(e^{h(x-t)} + e^{h(x+t)}) dx \\ &\geq \int_K e^{(h(x-t)+h(x+t))/2} dx. \end{aligned}$$

By Taylor's expansion we have for some $\theta \in [0, 1]$,

$$\frac{h(x-t) + h(x+t)}{2} = h(x) + \frac{1}{4}(\langle \text{Hess}h(x+\theta t)t, t \rangle + \langle \text{Hess}h(x-\theta t)t, t \rangle) \geq h(x) - \frac{1}{2}\alpha|t|^2.$$

Thus

$$\mu(t + K) \geq \int_K e^{h(x)-\alpha|t|^2/2} = e^{-\alpha|t|^2/2} \mu(K)$$

and

$$1 \geq \sum_{i=1}^N \mu(t_i + 3AT^0) \geq \sum_{i=1}^N e^{-\alpha|t_i|^2/2} \mu(3AT^0) \geq \frac{2N}{3} e^{-\alpha\delta^2/2} \geq N^{1/3} e^{-\alpha\delta^2/2}$$

and (9) easily follows. □

Proposition 14. *Suppose that X satisfies the L_2 -Sudakov minoration with constant C_X . Then for any $p \geq 2$*

$$N\left(\mathcal{M}_p(X), \frac{eC_X}{\sqrt{p}} B_2^n\right) \leq e^p.$$

In particular if X is isotropic we have for $2 \leq p \leq n$,

$$\left(\mathbb{E}\|X\|_{\mathcal{Z}_p(X)}^2\right)^{1/2} \leq e\left(C_X \sqrt{\frac{n}{p}} + 1\right).$$

Proof. Suppose that $N = N(\mathcal{M}_p(X), eC_X p^{-1/2} B_2^n) \geq e^p$. We can choose $t_1, \dots, t_N \in \mathcal{M}_p(X)$, such that $\|t_i - t_j\|_2 \geq eC_X p^{-1/2}$. We have

$$\mathbb{E} \sup_{i \geq N} \langle t_i, X \rangle \geq \frac{1}{C_X} eC_X p^{-1/2} \sqrt{\log N} > e.$$

However on the other hand,

$$\mathbb{E} \sup_{i \geq N} \langle t_i, X \rangle \leq \left(\mathbb{E} \sup_{i \geq N} |\langle t_i, X \rangle|^p \right)^{1/p} \leq \left(\sum_{i \geq N} \mathbb{E} |\langle t_i, X \rangle|^p \right)^{1/p} \leq N^{1/p} \max_i \|\langle t_i, X \rangle\|_p \leq e.$$

To show the second estimate we proceed in a similar way as in the proof of Proposition 8. We choose $T \subset \mathcal{M}_p(X)$ such that $|T| \leq e^p$ and $\mathcal{M}_p(X) \subset T + eC_X p^{-1/2} B_2^n$. We have

$$\|X\|_{\mathcal{Z}_p(X)} \leq eC_X p^{-1/2} |X| + \sup_{t \in T} |\langle t, X \rangle|$$

so that

$$\left(\mathbb{E} \|X\|_{\mathcal{Z}_p(X)}^2 \right)^{1/2} \leq eC_X p^{-1/2} (\mathbb{E} |X|^2)^{1/2} + \left(\mathbb{E} \sup_{t \in T} |\langle t, X \rangle|^2 \right)^{1/2}.$$

Vector X is isotropic, so $\mathbb{E} |X|^2 = n$ and since $T \subset \mathcal{M}_p(X)$ and $p \geq 2$ we get

$$\begin{aligned} \left(\mathbb{E} \sup_{t \in T} |\langle t, X \rangle|^2 \right)^{1/2} &\leq \left(\mathbb{E} \sup_{t \in T} |\langle t, X \rangle|^p \right)^{1/p} \leq \left(\sum_{t \in T} \mathbb{E} |\langle t, X \rangle|^p \right)^{1/p} \\ &\leq |T|^{1/p} \max_{t \in T} \|\langle t, X \rangle\|_p \leq e. \end{aligned}$$

□

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