

Statistical phylogenetics

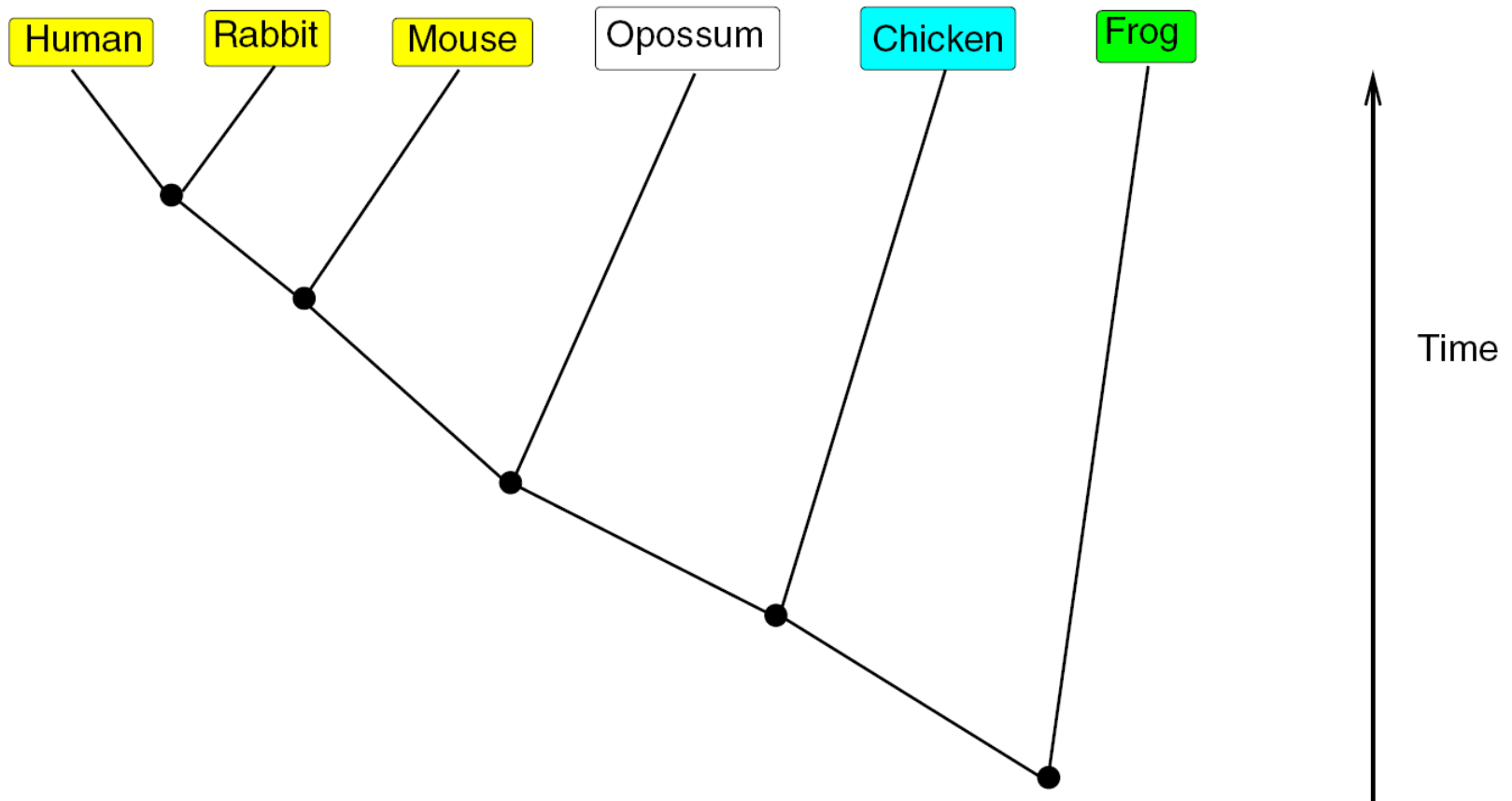
Niko Beerenwinkel



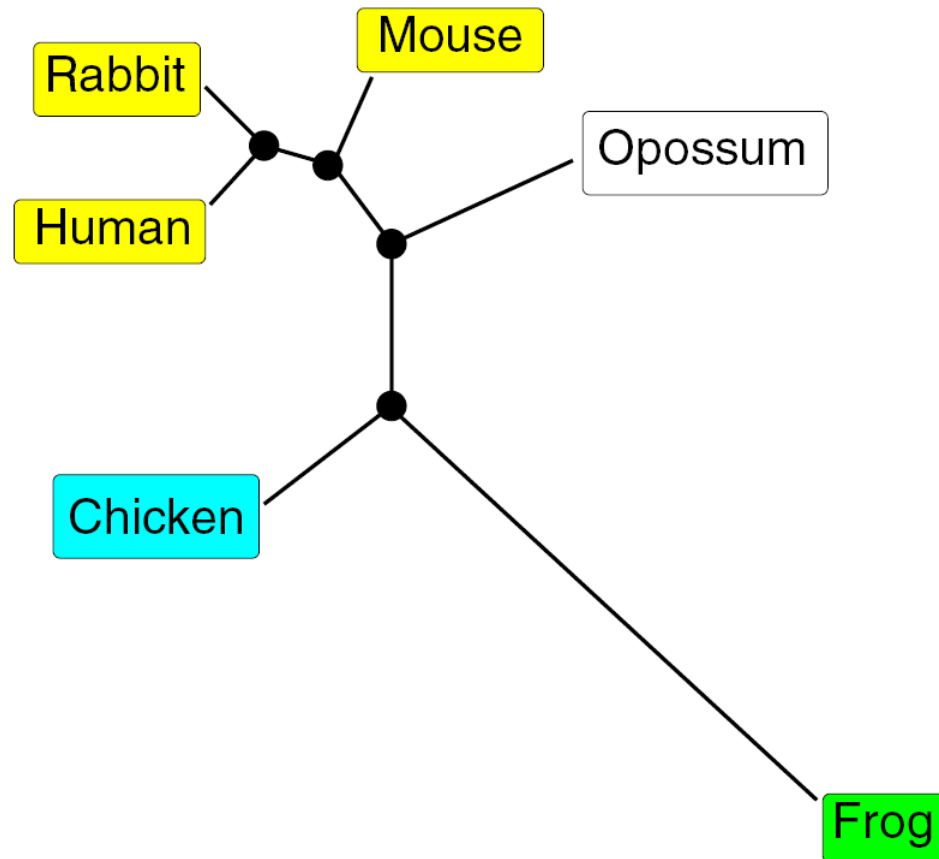
Outline

- Phylogenetic trees
- Nucleotide substitution models
- Likelihood
- ML estimation and bootstrapping
- Bayesian inference
- Rate heterogeneity
- Phylo-HMMs

Rooted phylogenetic tree



Unrooted phylogenetic tree



Phylogenetic trees

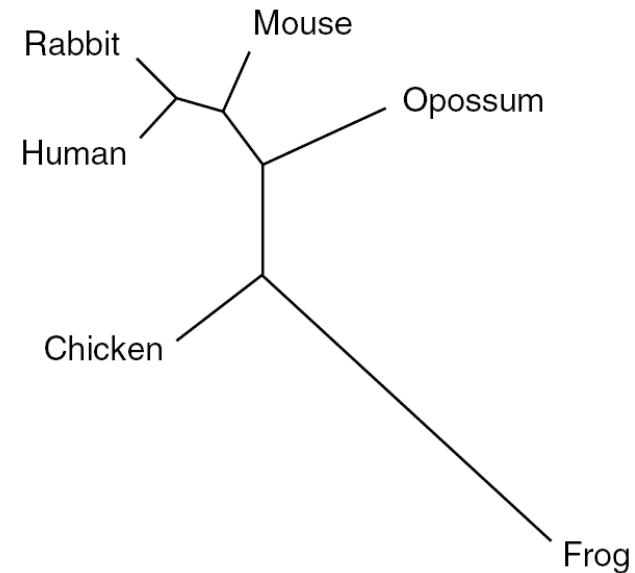
- Leaves = contemporary species
- Interior vertices = common ancestors
- Topology (graph structure) defines branching order (subtrees)
- Branch lengths (parameters) define time
 - absolute time t , or
 - phylogenetic time $w = \lambda t$ (λ = nucleotide substitution rate)

Phylogenetic inference

- Given a multiple alignment

Frog	G	C	T	T	G	A	C	T	T	C	T	G	A	G	G	T	T
Chicken	G	C	G	T	A	A	C	T	T	C	A	C	A	T	G	A	T
Human	G	C	G	T	C	A	C	T	T	G	A	G	A	C	G	C	T
Rabbit	G	C	G	T	C	A	C	T	T	G	A	G	A	C	G	C	T
Mouse	G	C	G	T	C	A	C	T	T	G	A	C	A	G	G	C	T
Opossum	G	C	G	T	C	A	C	T	T	G	A	G	A	C	G	C	T

- Find the best tree explaining the data

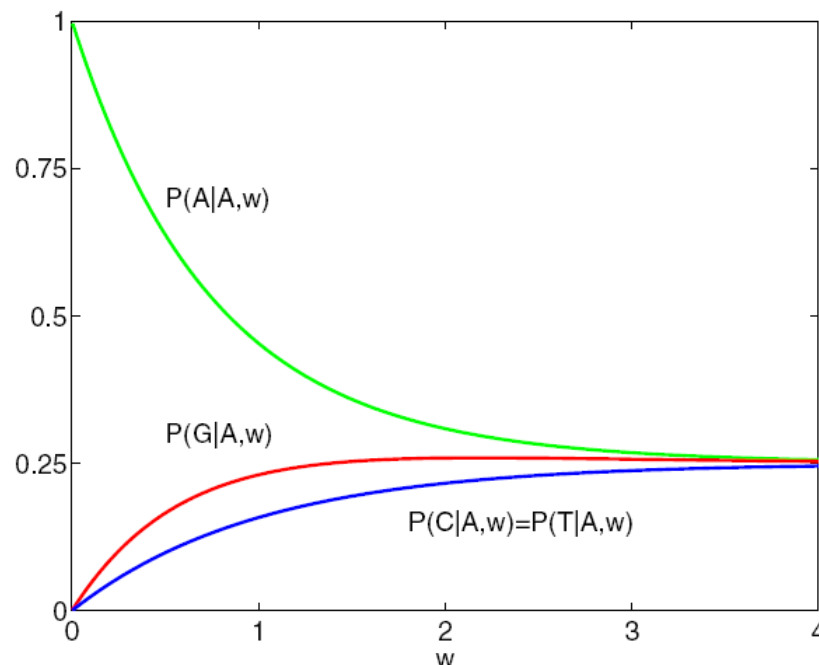
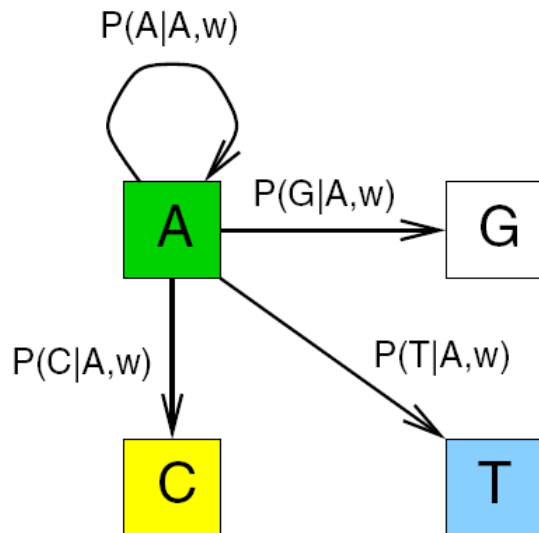


Phylogeny reconstruction methods

- Distance-based clustering methods
 - Define an evolutionary distance
 - Use hierarchical clustering
 - UPGMA, Neighbor joining
 - Shortcomings: information loss, not robust against violations of tree assumption
- Parsimony
 - “Minimum evolution” principle: Find the tree explaining the data by the minimum number of mutations.
 - Shortcomings: “model-free”, not consistent
- Likelihood methods

Nucleotide substitution models

- For characters $x, y \in \{A, C, G, T\}$, we define $P(y | x, w)$ as the probability of observing y after w time units given that the same site was originally occupied by x .



Continuous-time homogeneous Markov chain

- Let $y_i(t) \in \{A, C, G, T\}$ be the nucleotide at position i at time t .
- We assume for all $s, t > 0$ and $i, k \in \{1, \dots, N\}$,

- 1) a Markov process

$$P[y_i(t + \Delta t) | y_i(t), y_i(t - \Delta t), \dots] = P[y_i(t + \Delta t) | y_i(t)]$$

- 2) homogeneous in time

$$P[y_i(s + t) | y_i(s)] = P[y_i(t) | y_i(0)]$$

- 3) identical across sites

$$P[y_i(t) | y_i(0)] = P[y_k(t) | y_k(0)]$$

- 4) independent among sites

$$P[y_1(t), \dots, y_N(t) | y_1(0), \dots, y_N(0)] = \prod_{i=1}^N P[y_i(t) | y_i(0)]$$

Transition matrix

- The nucleotide substitution process is defined by the 4×4 transition matrix

$$\mathbf{P}(t) = \left(P[y(t) = a \mid y(0) = b] \right)_{a,b \in \{A,C,G,T\}}$$

- $\mathbf{P}(0) = \mathbf{I}$
- Chapman-Kolmogorov equation for continuous-time homogeneous Markov chains:

$$\mathbf{P}(t + s) = \mathbf{P}(t)\mathbf{P}(s)$$

for all $s, t > 0$.

Rate matrix

- Ansatz: $\mathbf{P}(dt) = \mathbf{P}(0) + \mathbf{R}dt$, where \mathbf{R} is the *rate matrix*.
- Then

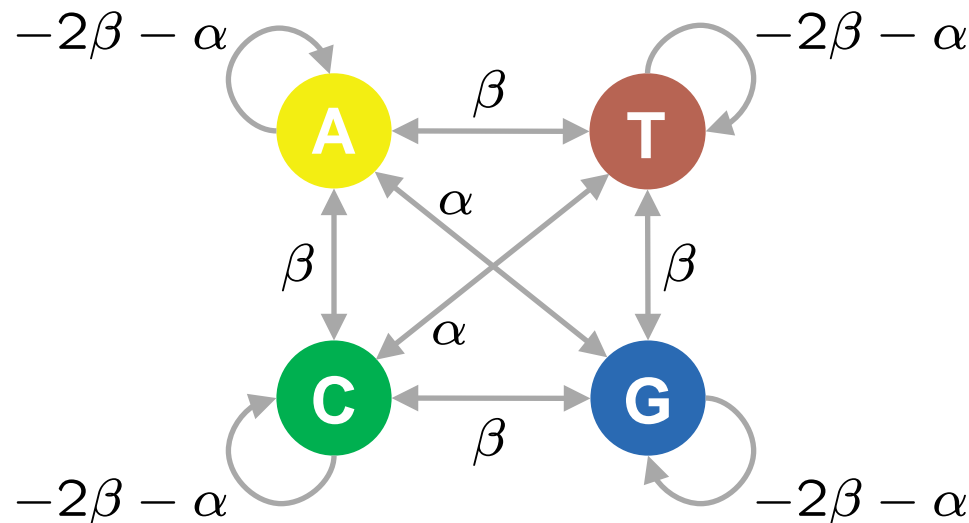
$$\mathbf{P}(t + dt) = \mathbf{P}(dt)\mathbf{P}(t) = (\mathbf{I} + \mathbf{R}dt)\mathbf{P}(t)$$

$$\Rightarrow \frac{d\mathbf{P}(t)}{dt} = \mathbf{R}\mathbf{P}(t)$$

$$\Rightarrow \mathbf{P}(t) = \exp(\mathbf{R}t) = \sum_{k=0}^{\infty} \frac{1}{k!} (\mathbf{R}t)^k$$

Example: Kimura model

- Purines = {A, G}
- Pyrimidines = {C, T}
- Transitions (rate α)
 - purine \leftrightarrow purine
 - pyrimidine \leftrightarrow pyrimidine
- Transversions (rate β)
 - purine \leftrightarrow pyrimidine



$$\mathbf{R} = \begin{pmatrix} -2\beta - \alpha & \beta & \alpha & \beta \\ \beta & -2\beta - \alpha & \beta & \alpha \\ \alpha & \beta & -2\beta - \alpha & \beta \\ \beta & \alpha & \beta & -2\beta - \alpha \end{pmatrix}$$

Equilibrium base distribution

- The marginal distribution of nucleotides

$$\mathbf{u}(w) = (P[y(w) = a])_{a \in \{A,C,G,T\}}$$

defines a homogeneous Markov chain,

$$\mathbf{u}(v + w) = \mathbf{P}(w)\mathbf{u}(v)$$

- An ergodic Markov chain converges to a unique stationary distribution

$$\lim_{w \rightarrow \infty} \mathbf{u}(w) = \pi = (\pi_A, \pi_C, \pi_G, \pi_T)$$

characterized by $\mathbf{P}(w)\pi = \pi$.

- For the Kimura model, we find $\pi = \left(\frac{1}{4}, \frac{1}{4}, \frac{1}{4}, \frac{1}{4}\right)$.

HKY85 model

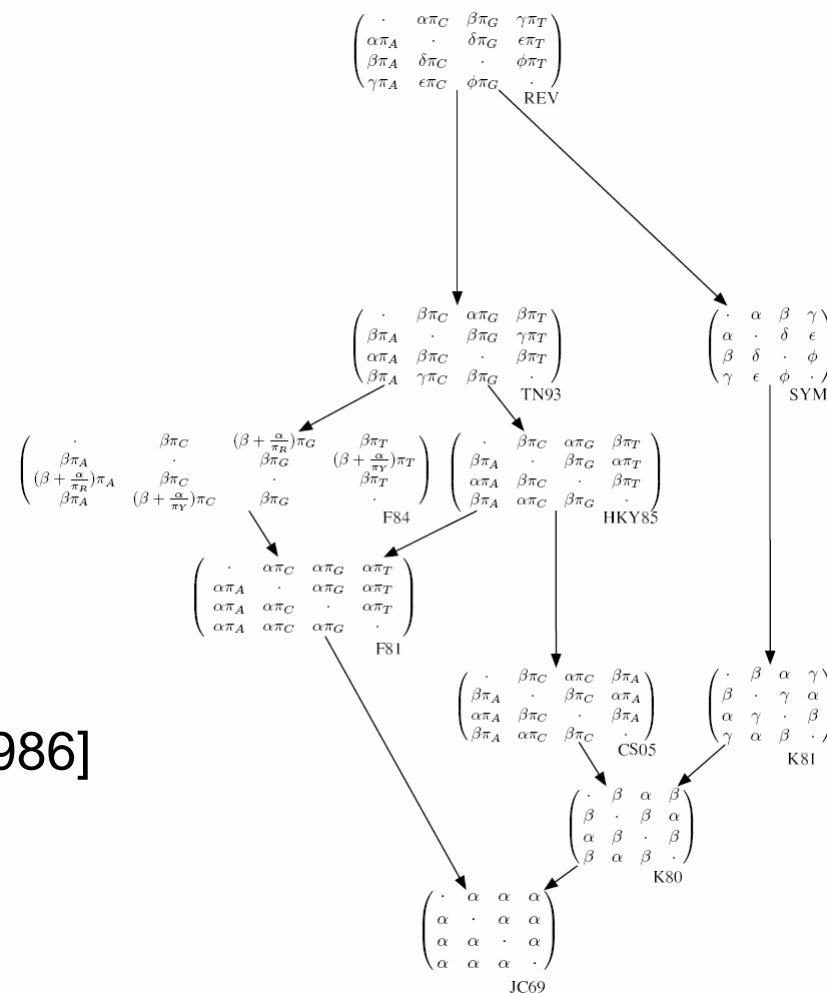
- The following modified rate matrix has the stationary distribution $\pi = (\Pi_A, \Pi_C, \Pi_G, \Pi_T)$

$$\mathbf{R} = \begin{pmatrix} * & \Pi_A \beta & \Pi_A \alpha & \Pi_A \beta \\ \Pi_C \beta & * & \Pi_C \beta & \Pi_C \alpha \\ \Pi_G \alpha & \Pi_G \beta & * & \Pi_G \beta \\ \Pi_T \beta & \Pi_T \alpha & \Pi_T \beta & * \end{pmatrix}$$

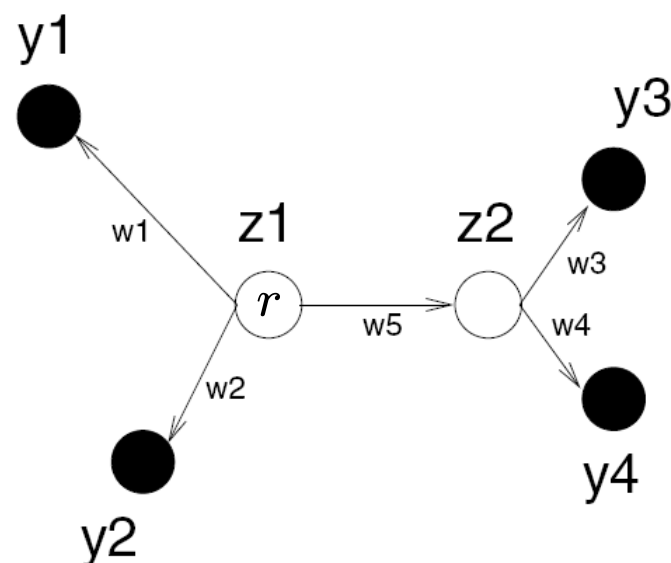
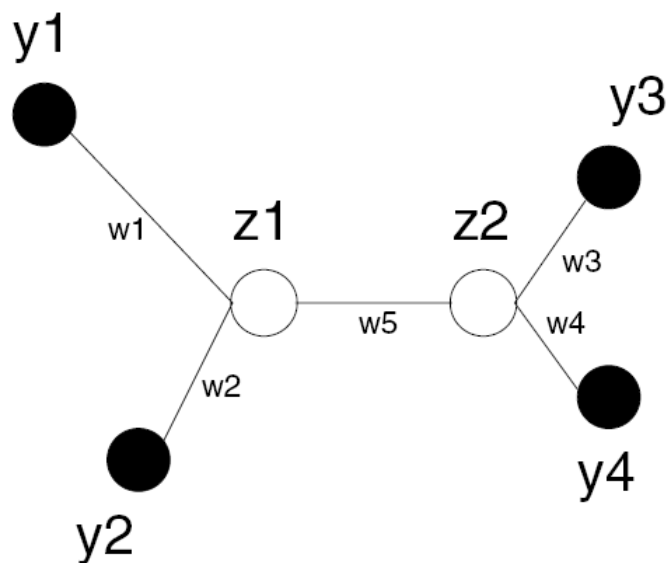
- We assume *stationarity* of the Markov chain: The nucleotide distribution is equal to π over the whole tree.

Felsenstein hierarchy

- JC69 [Jukes and Cantor, 1969]
- K80 [Kimura, 1980]
- K81 [Kimura, 1981]
- CS05 [Yap and Pachter, 2004]
- F81 [Felsenstein, 1981]
- HKY85 [Hasegawa et al., 1985]
- F84 [Felsenstein, 1989]
- TN93 [Tamura and Nei, 1993]
- SYM [Zharkikh, 1994]
- REV [Lanave et al., 1984, Tavaré, 1986]



A phylogenetic tree is a Bayesian network



$$\begin{aligned}
 P(y_1, y_2, y_3, y_4, z_1, z_2 \mid \mathbf{w}) &= \prod(z_1) P(y_1 \mid z_1, w_1) P(y_2 \mid z_1, w_2) \cdot \\
 &\cdot P(z_2 \mid z_1, w_5) \cdot \\
 &\cdot P(y_3 \mid z_2, w_3) P(y_4 \mid z_2, w_4)
 \end{aligned}$$

The nucleotide substitution model defines the LPDs

- In general, $x_i \in \{A, C, G, T\}$ is the random variable indicating the nucleotide at vertex i .

$$\begin{aligned} P(x_1, \dots, x_M) &= \prod(x_r) \prod_{i \in V \setminus \{r\}} P(x_i | x_{\text{pa}(i)}, w_i) \\ &= \prod(x_r) \prod_{i \in V \setminus \{r\}} \mathbf{P}(w_i)_{x_i, x_{\text{pa}(i)}} \end{aligned}$$

- Choice of the root vertex does not matter as long as the Markov chain is *reversible*, i.e.,

$$P(y | x, w) \prod(x) = P(x | y, w) \prod(y)$$

Marginalization over extinct species

- \mathbf{y} = extant (contemporary) species
- \mathbf{z} = extinct common ancestors

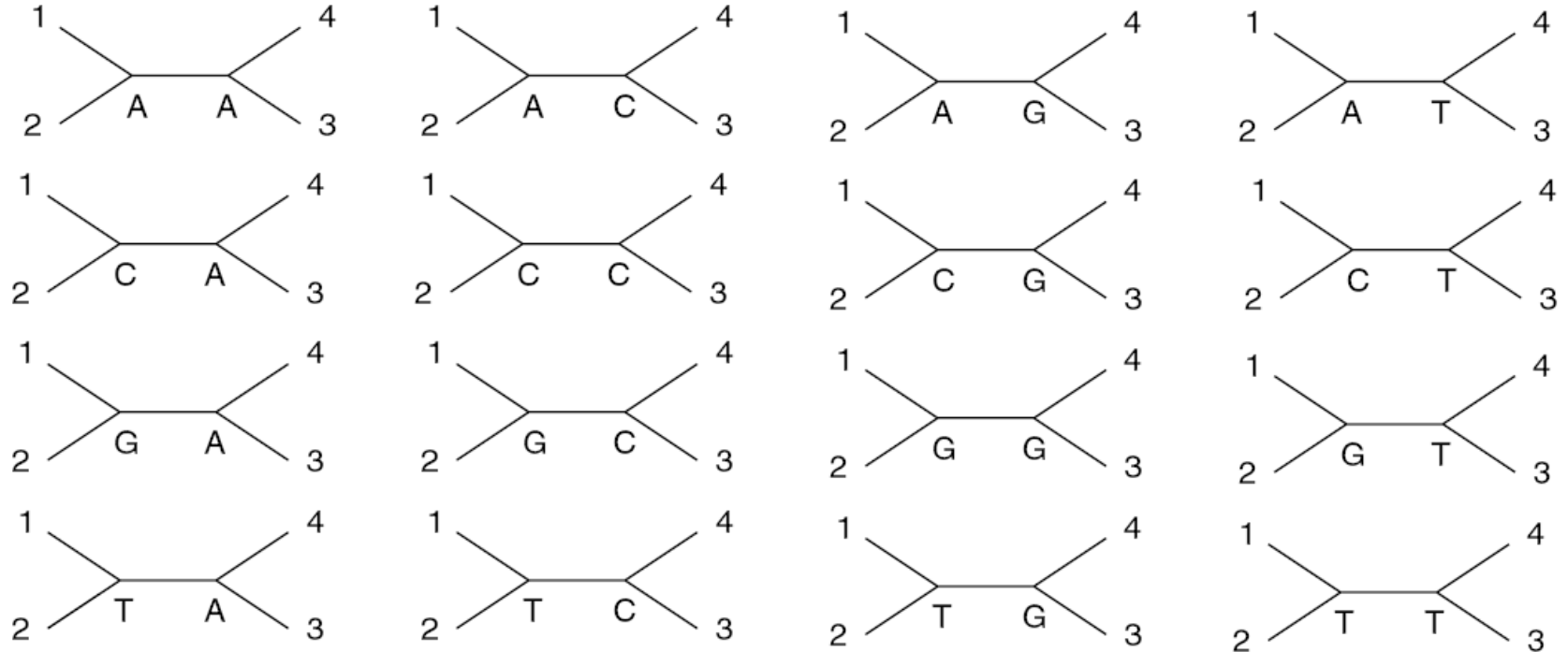
$$P(\mathbf{y} \mid \mathbf{w}, S) = \sum_{\mathbf{z}} P(\mathbf{y}, \mathbf{z} \mid \mathbf{w}, S)$$

where S indicates the tree topology.

- This marginal distribution can be computed efficiently with the sum-product algorithm (\rightarrow “peeling algorithm”, “Felsenstein algorithm”), a generalization of the forward algorithm.

Example

- Two extinct species (= hidden variables)



Likelihood of a phylogenetic tree

- Given a multiple alignment $\mathcal{D} = \{\mathbf{y}_1, \dots, \mathbf{y}_N\}$, where \mathbf{y}_t is the alignment column at position t , the likelihood of tree topology S and branch lengths \mathbf{w} is

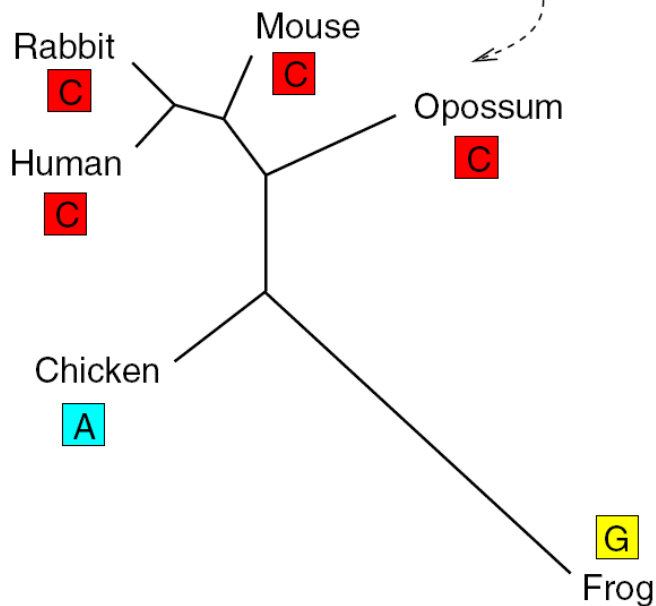
$$\begin{aligned} P(\mathcal{D} \mid \mathbf{w}, S) &= \prod_{t=1}^N P(\mathbf{y}_t \mid \mathbf{w}, S) \\ &= \prod_{t=1}^N \sum_{\mathbf{z}} P(\mathbf{y}_t, \mathbf{z}_t \mid \mathbf{w}, S) \end{aligned}$$

- We have omitted here and will continue to omit the parameters of the nucleotide substitution model (the rate matrix).

Likelihood of a phylogenetic tree

↓

Frog	G	C	T	T	G	A	C	T	T	C	T	G	A	G	G	T	T
Chicken	G	C	G	T	A	A	C	T	T	C	A	C	A	T	G	A	T
Human	G	C	G	T	C	A	C	T	T	G	A	G	A	C	G	C	T
Rabbit	G	C	G	T	C	A	C	T	T	G	A	G	A	C	G	C	T
Mouse	G	C	G	T	C	A	C	T	T	G	A	C	A	G	G	C	T
Opossum	G	C	G	T	C	A	C	T	T	G	A	G	A	C	G	C	T



Maximum likelihood

- No analytical solution exists for the MLE problem

$$\max_{S, \mathbf{w}} P(\mathcal{D} \mid \mathbf{w}, S)$$

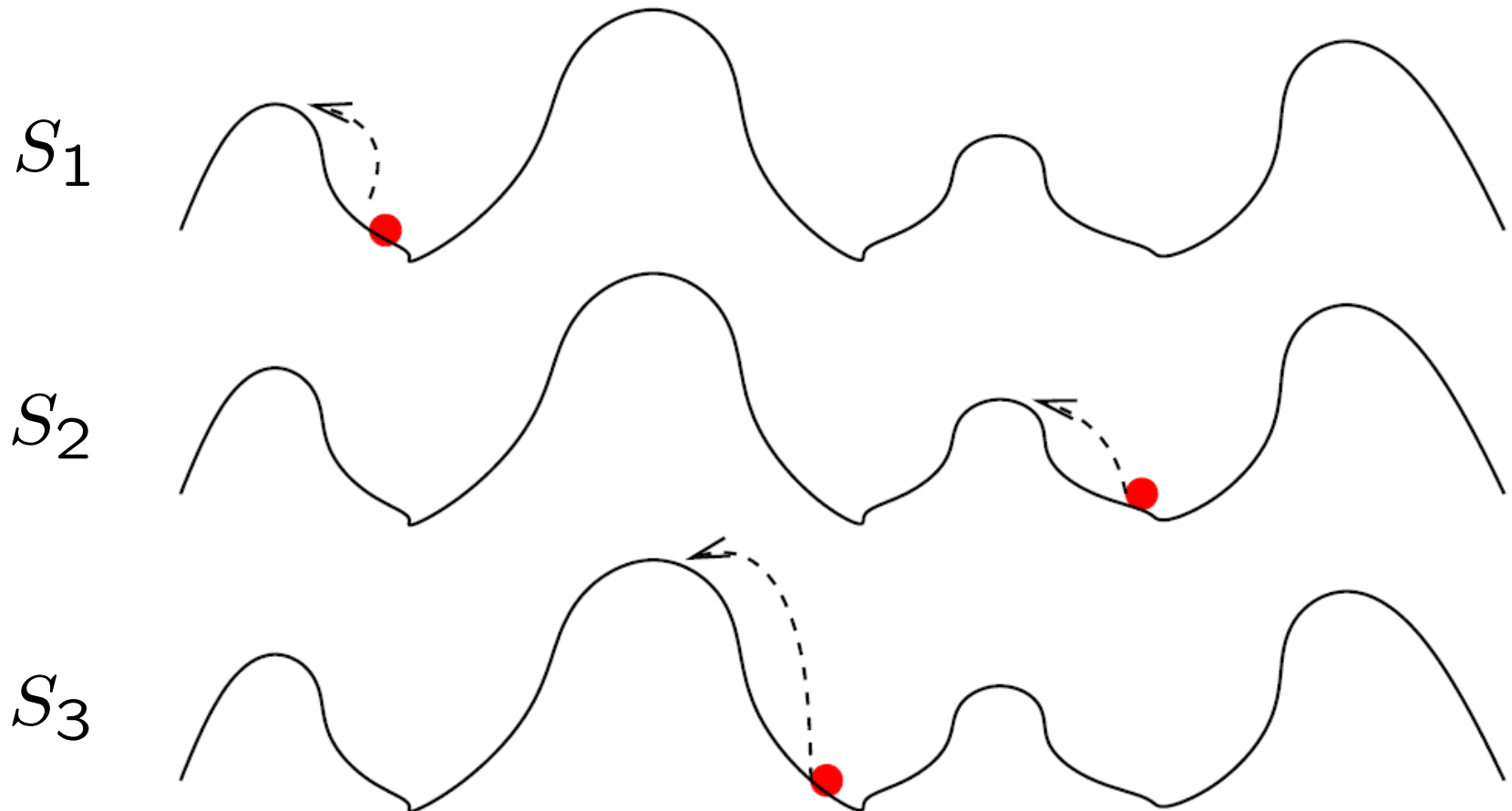
and numerical optimization is NP-hard.

- Branch lengths are optimized by a gradient ascent scheme:

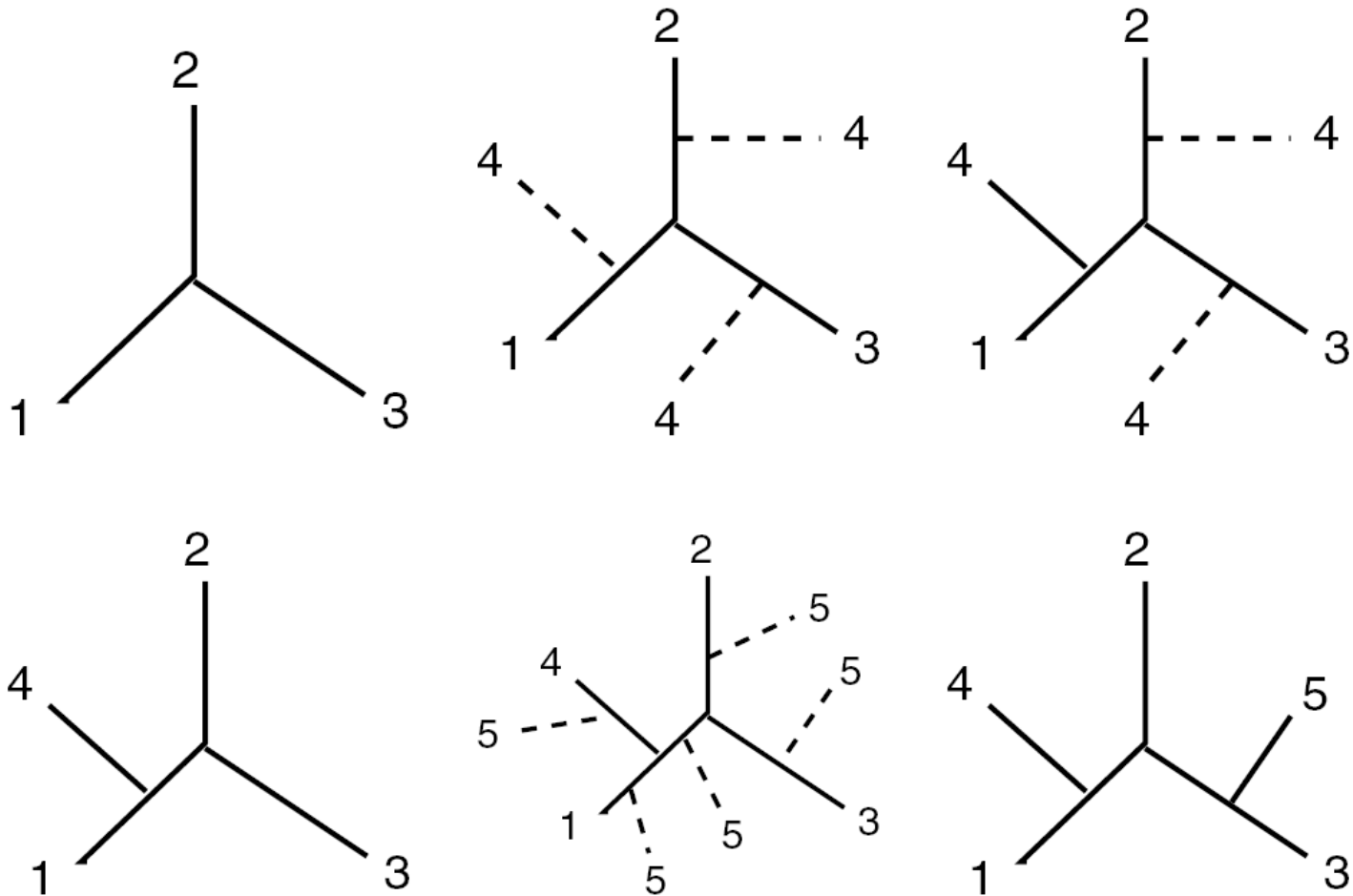
$$\mathbf{w} \rightarrow \mathbf{w} + \mathbf{A} \nabla_{\mathbf{w}} \log P(\mathcal{D} \mid \mathbf{w}, S)$$

- There are $(2n - 5)!!$ unrooted tree topologies for n taxa.
⇒ heuristic search procedures:
 - DNAML
 - Quartet puzzling

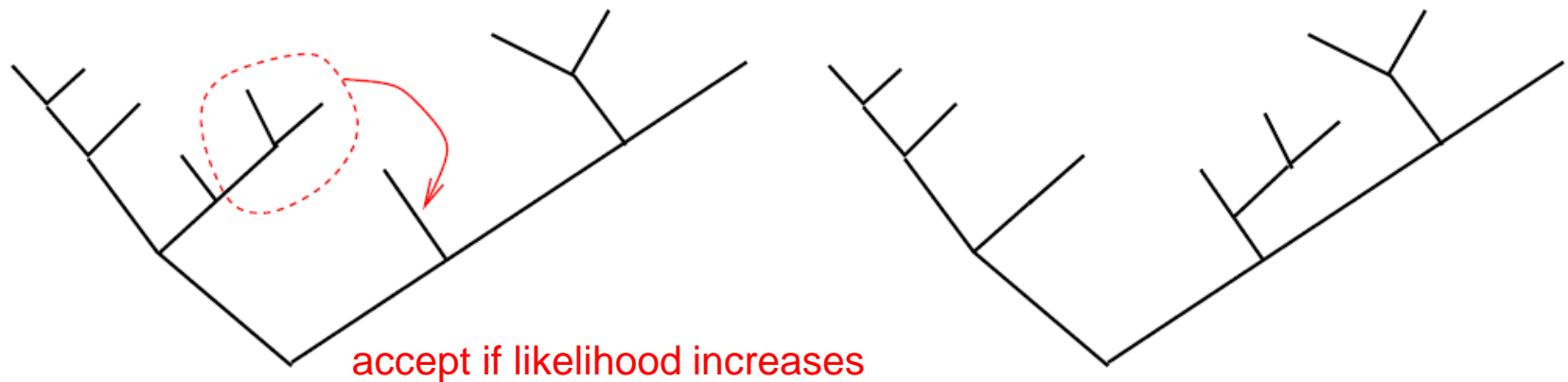
Branch lengths are optimized for each topology



DNAML: Iterative attachment of branches



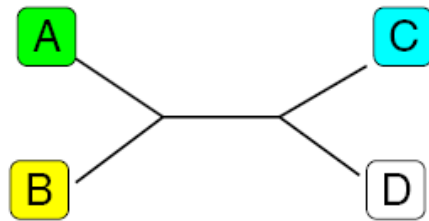
DNAML: Branch regrafting



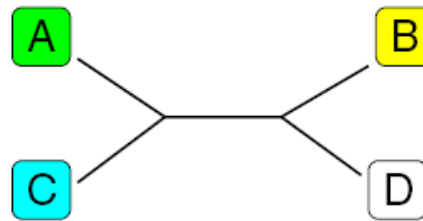
- DNAML employs greedy search strategies
- Results depend on the order in which alignment columns are considered
- Only few branch manipulations can be computationally afforded

Quartet puzzling

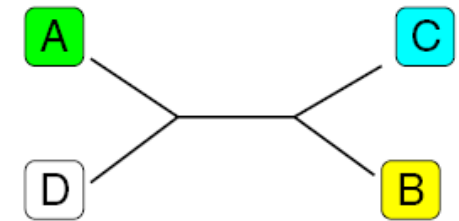
1. Construct all (n choose 4) quartet trees on four of the n given taxa using maximum likelihood



Topology 1



Topology 2

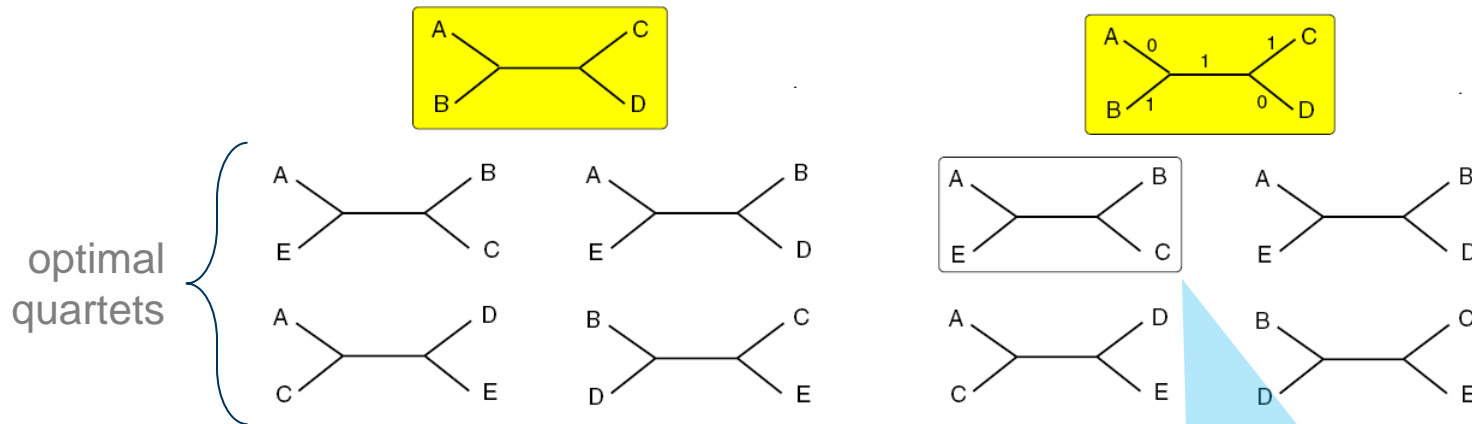


Topology 3

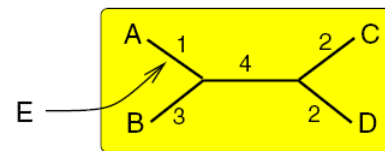
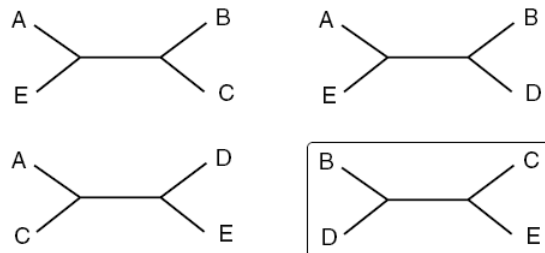
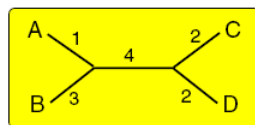
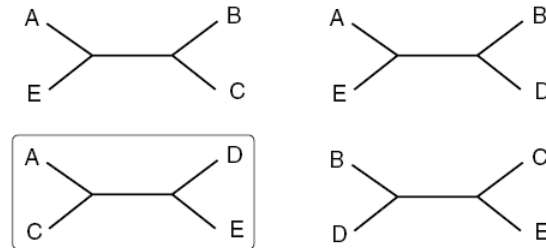
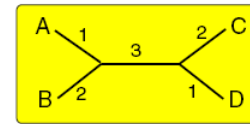
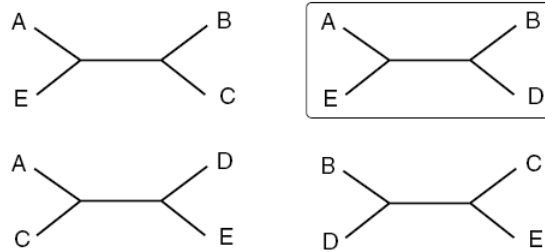
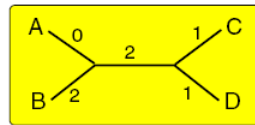
2. Combine the quartet trees into a global tree by adding taxa iteratively minimizing conflicts (puzzling step).

Example

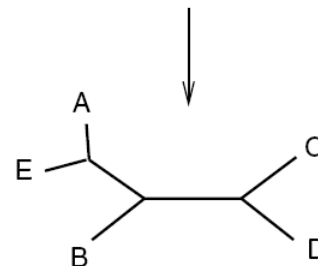
- Consider five taxa A, B, C, D, and E.
- There are $(5 \text{ choose } 4) = 5$ quartets.



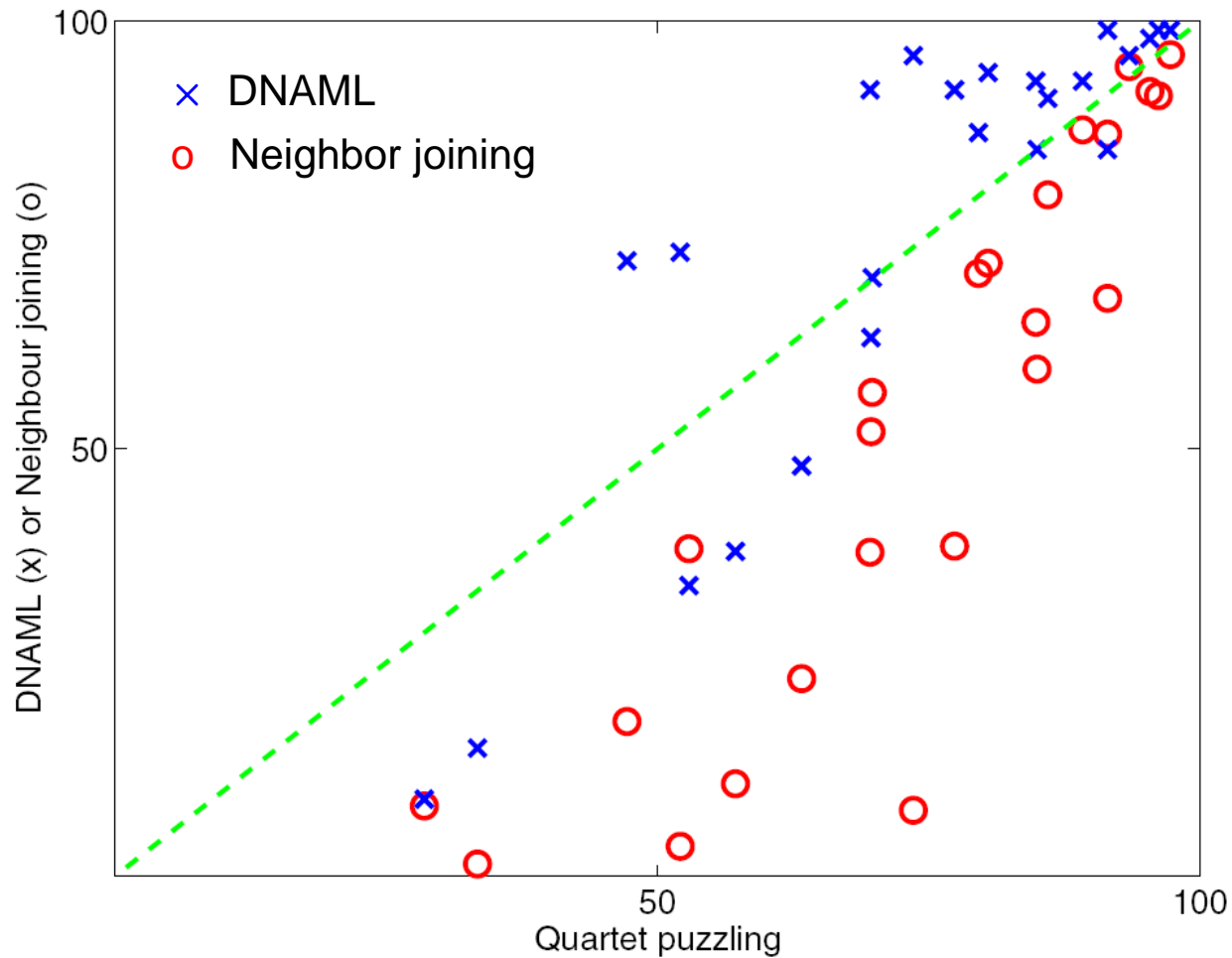
Example



select least penalized branch

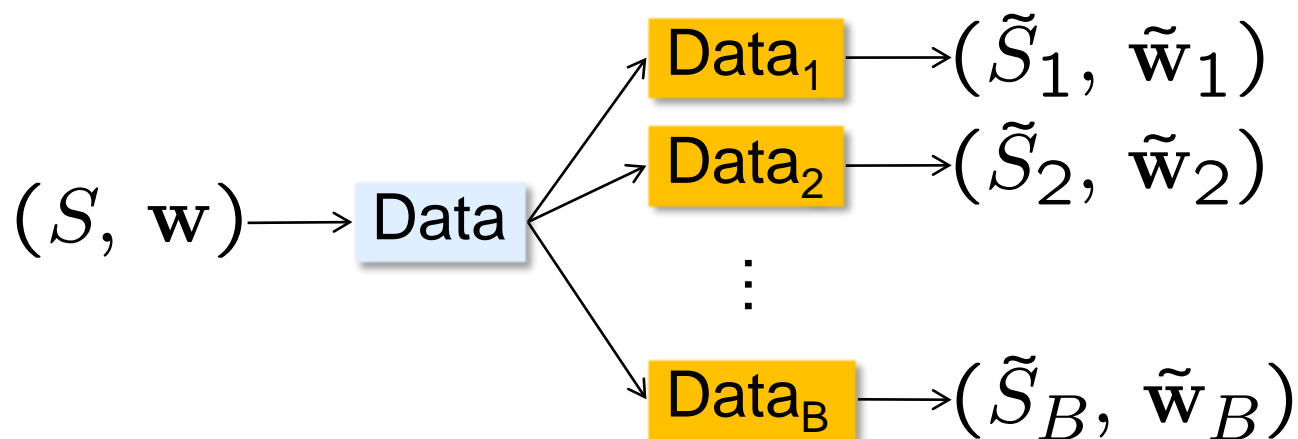


Performance of quartet puzzling



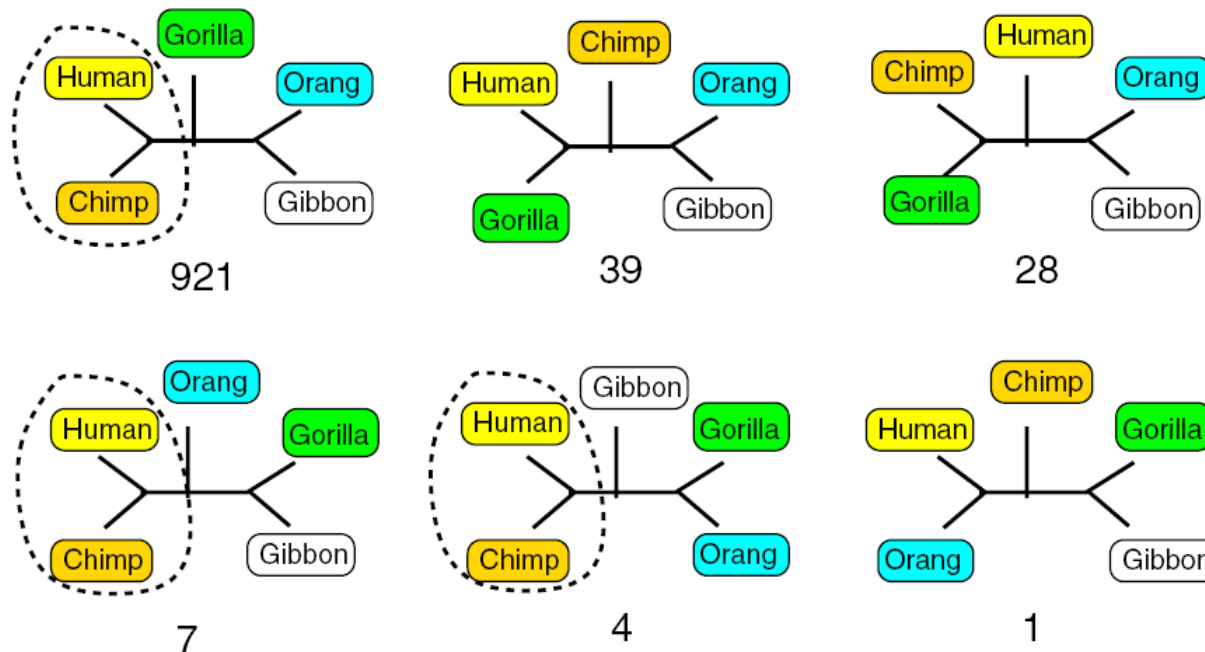
Bootstrapping phylogenetic trees

- How sure can we be that the estimated ML phylogenetic tree is correct?



Example: (Human, Chimp)

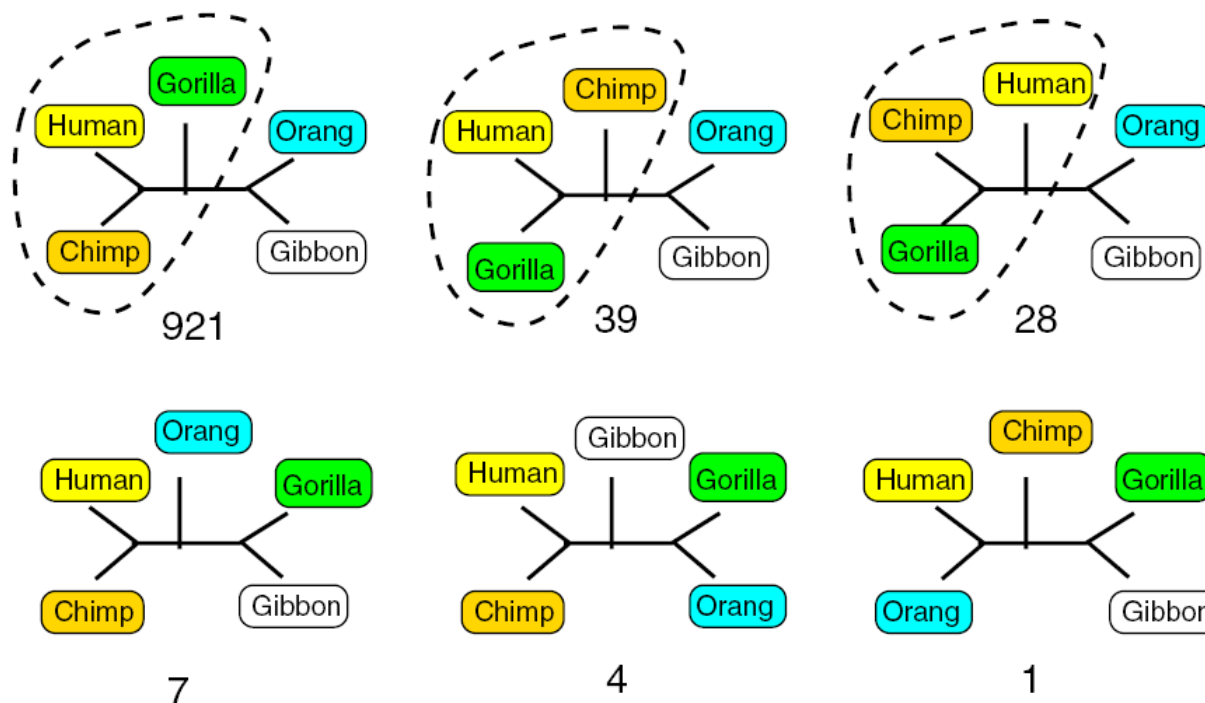
- $B = 1000$ bootstrap samples



$$P[(\text{Human, Chimp})] = \frac{921 + 7 + 4}{1000} = 0.932$$

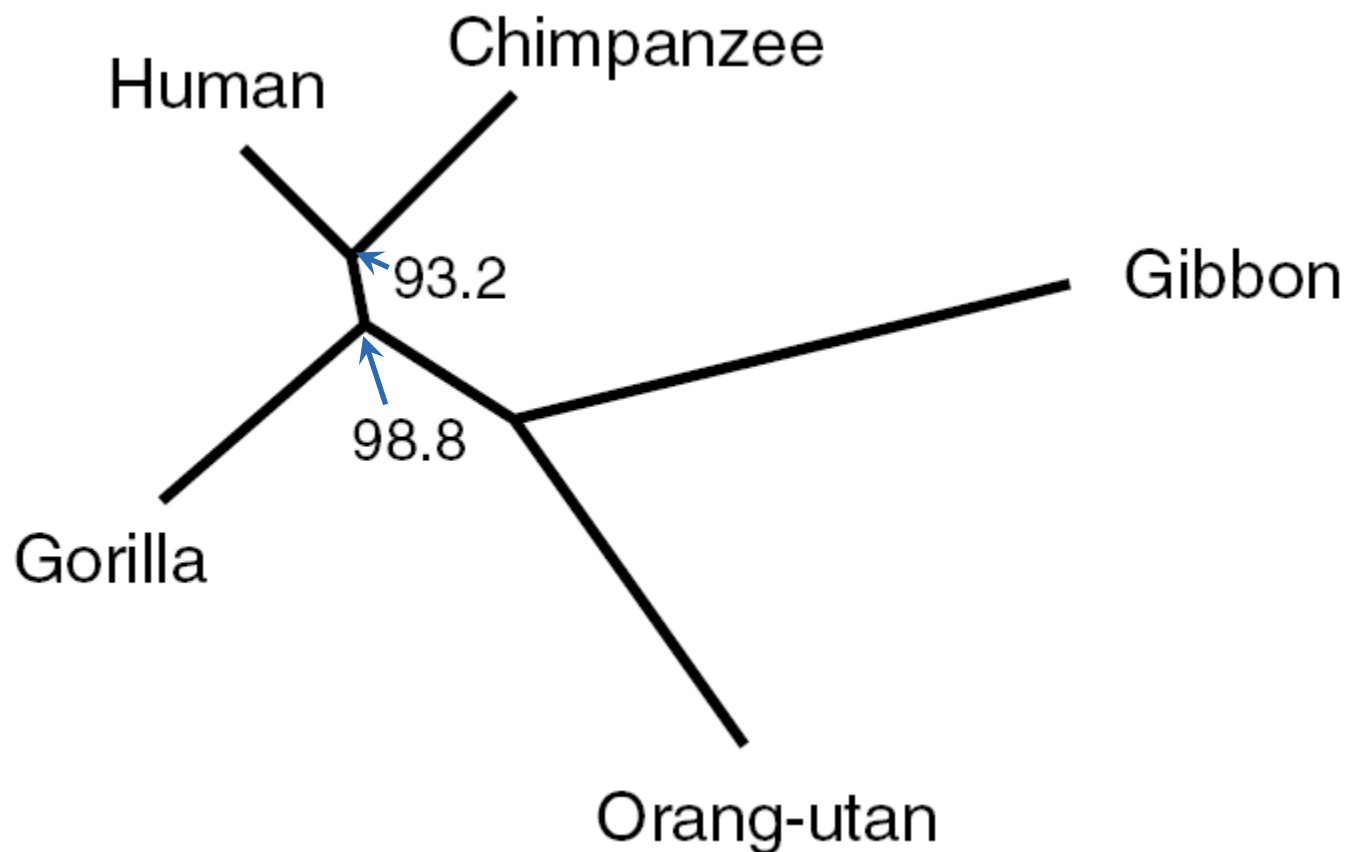
Example: (Human, Chimp, Gorilla)

- Same B = 1000 bootstrap samples



$$P[\text{((Human, Chimp), Gorilla)}] = \frac{921 + 39 + 28}{1000} = 0.988$$

Bootstrapped tree

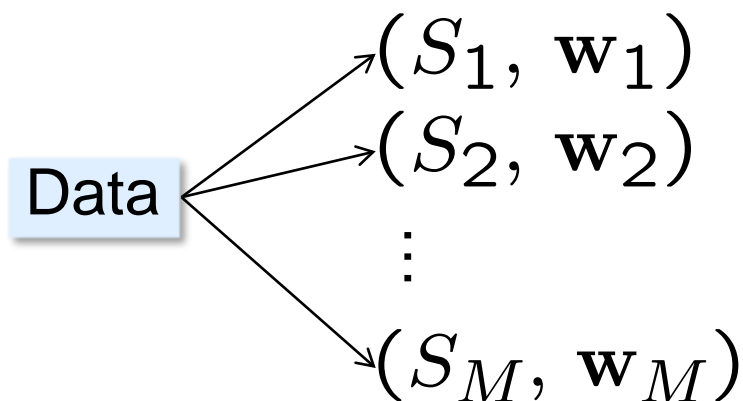


Bayesian inference

$$\begin{aligned} P(S | \mathcal{D}) &\propto P(S) P(\mathcal{D} | S) \\ &= P(S) \int P(\mathcal{D} | \mathbf{w}, S) P(\mathbf{w} | S) d\mathbf{w} \end{aligned}$$

but the marginal likelihood $P(\mathcal{D} | S)$ is analytically intractable

Sample from (S, w) and marginalize



$$P(S | \mathcal{D}) = \int P(S, \mathbf{w} | \mathcal{D}) d\mathbf{w}$$
$$\approx (\# \text{ trees with topology } S) / M$$

Rate heterogeneity

- Nucleotide substitution rates may vary across sites because of varying selective pressures. For example,
 - between coding and non-coding regions
 - among different regions of a protein (loops, catalytic residues)
 - among the three bases of a triplet coding for an amino acid
- Let us assume site-specific substitution rates r_t such that the local probabilities become $P(\mathbf{y}_t | r_t \mathbf{w}, S)$ and

$$P(\mathcal{D} | \mathbf{w}, S) = \int \prod_{t=1}^N P(\mathbf{y}_t | r_t \mathbf{w}, S) dP(r_1, \dots, r_N)$$

Independent substitution rates

- Let us assume that, for a hyperparameter v ,

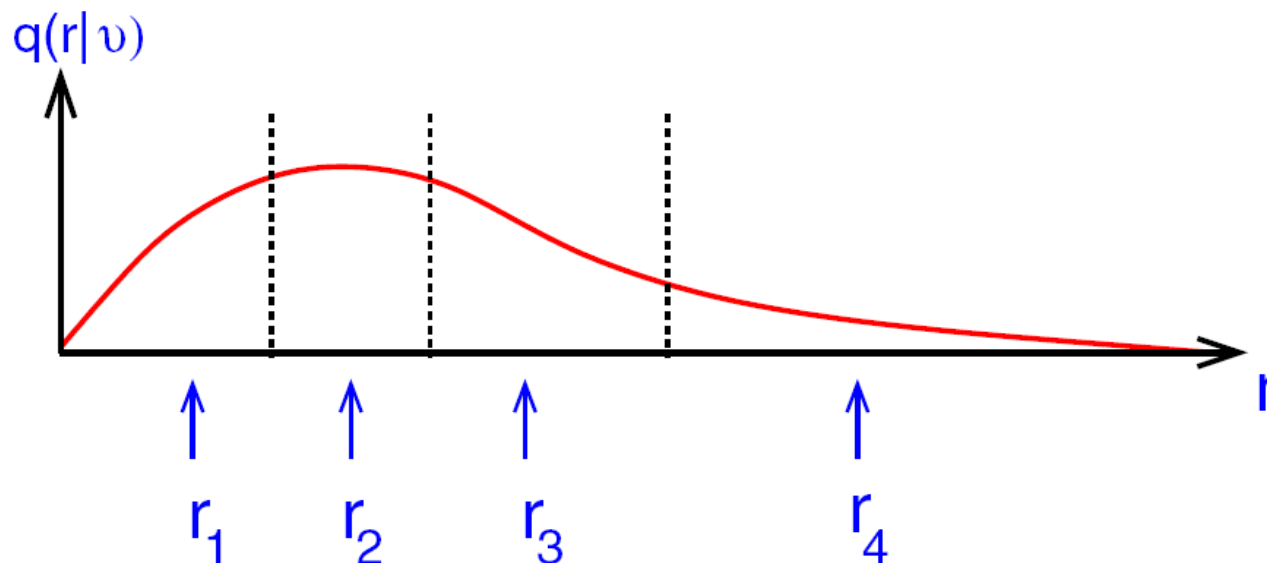
$$P(r_1, \dots, r_N) = \prod_{t=1}^N q(r_t | v)$$

- Then the likelihood simplifies to

$$P(\mathcal{D} | \mathbf{w}, v, S) = \prod_{t=1}^N \int_0^{\infty} P(\mathbf{y}_t | r_t \mathbf{w}, S) q(r_t | v) dr_t$$

- $q(r_t | v) = \Gamma(v, 1/v)$, gamma distribution with mean 1 and variance $1/v$.

Discrete gamma distribution



- Divide the positive real line into K intervals such that each interval contains an equal area of the gamma distribution.

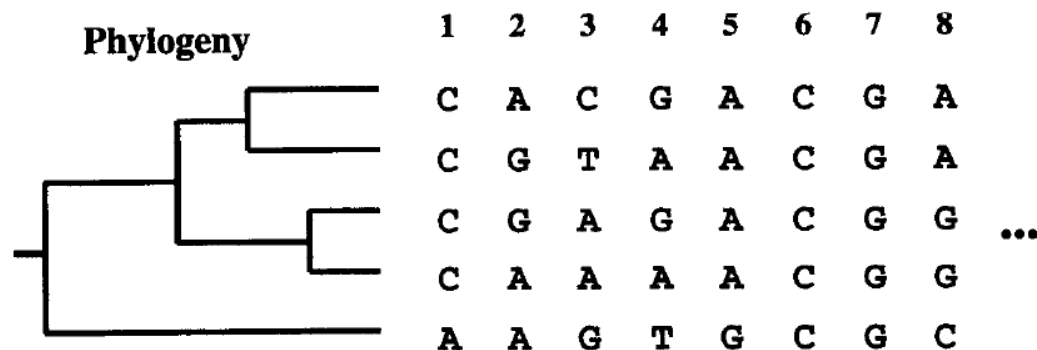
$$P(\mathcal{D} \mid \mathbf{w}, v, S) = \frac{1}{K} \prod_{t=1}^N \sum_{k=1}^K P(\mathbf{y}_t \mid r_k^v \mathbf{w}, S)$$

Spatial correlation among substitution rates

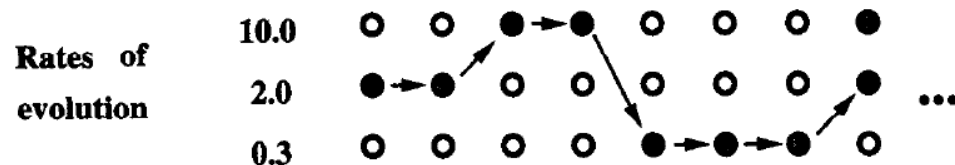
- Substitution rates will tend to be similar at neighboring sites.

- Setting
$$P(r_1, \dots, r_N) = q(r_1) \prod_{t=2}^N q(r_t | r_{t-1}, v)$$

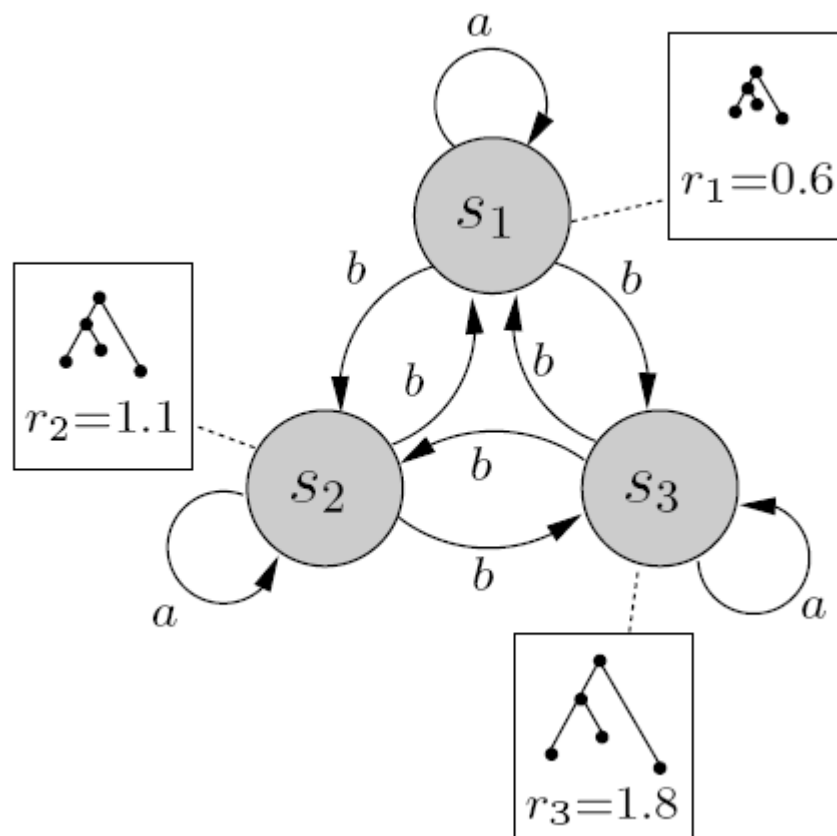
gives rise to a hidden Markov model:



Hidden Markov chain:

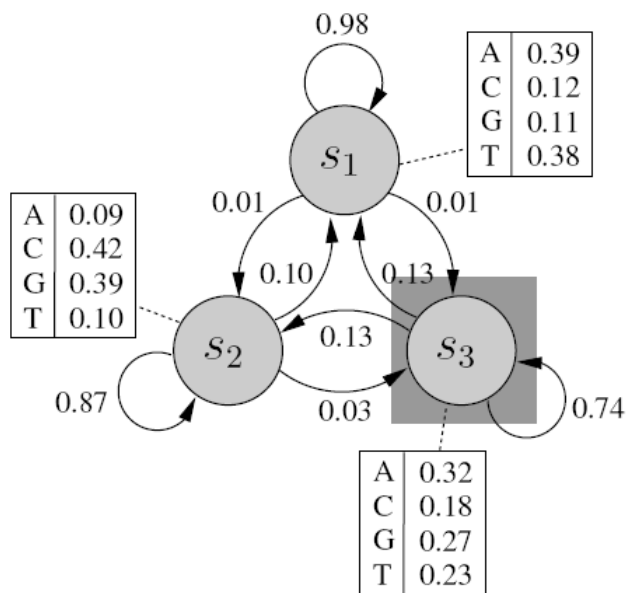


Markov chain state space



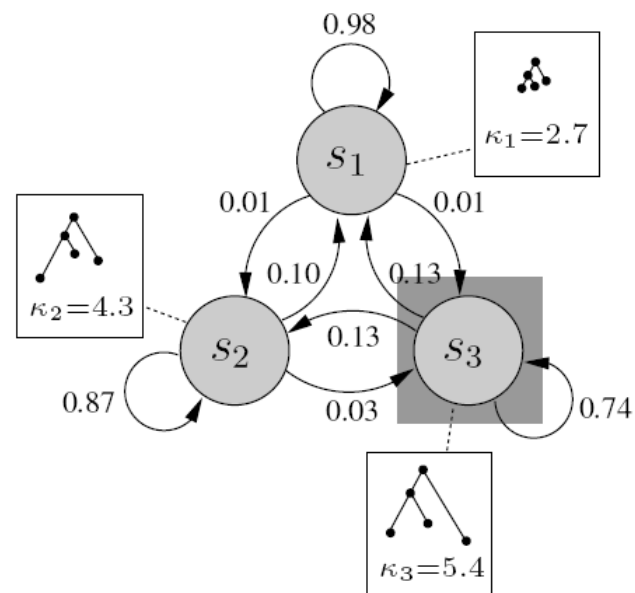
Phylogenetic hidden Markov models

HMM



$\mathbf{X} = \text{TAACGGCAGA} \dots$

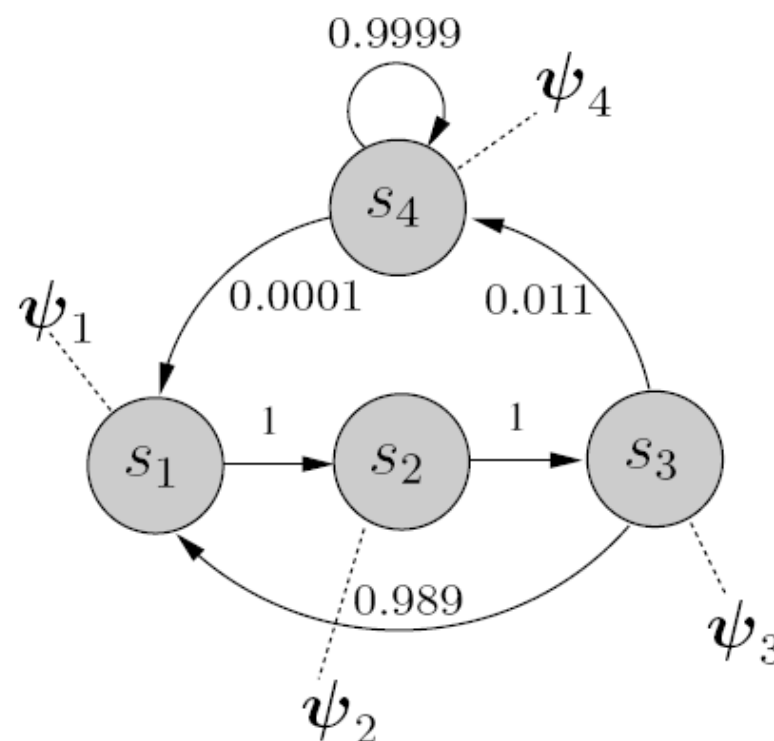
Phylo-HMM



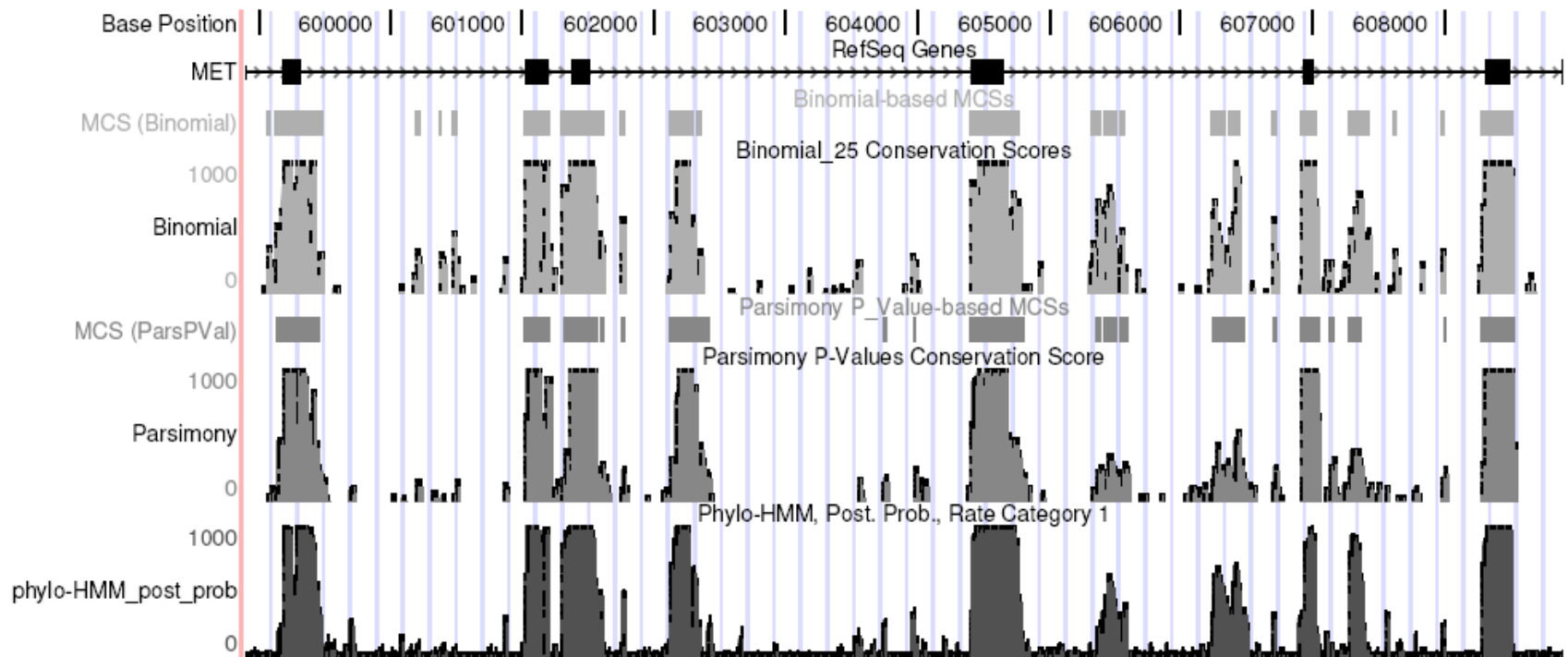
$\mathbf{X} = \begin{matrix} \text{TAACGGCAGA} \dots \\ \text{TTAGGCAAGG} \dots \\ \text{AAGGCGCCGA} \dots \end{matrix} \dots$

Phylo-HMM for gene finding

- Non-coding regions tend to have a higher substitution rate and a higher transition-transversion ratio
 - s_1, s_2, s_3 model codons
 - s_4 models non-coding sites
- ψ_1, \dots, ψ_4 capture the parameters of the phylogenetic models



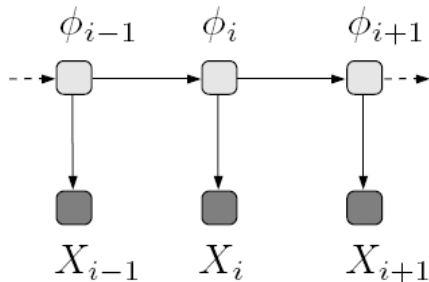
Detecting conserved genomic regions



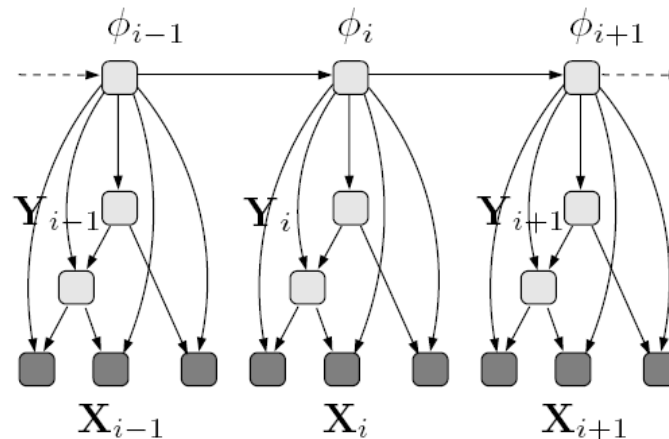
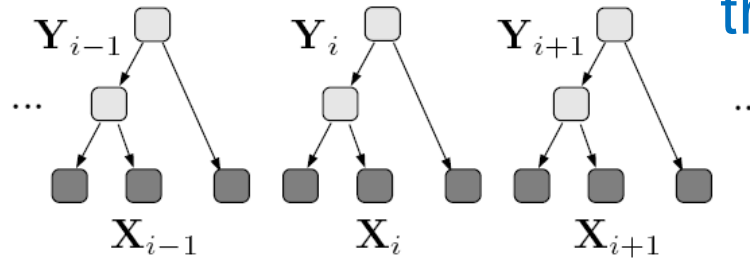
<http://genome.ucsc.edu>

Phylo-HMMs are Bayesian networks

HMM



Phylogenetic
tree model



Phylo-HMM

Varying tree topologies are also possible (detection of recombination).

Summary

- Probabilistic phylogenetic tree models are Bayesian networks with observed and hidden random variables.
- The LPDs are defined by nucleotide substitution models, which are examples of continuous-time Markov chains.
- Phylogenetic trees can be learned using ML or Bayes.
- Rate heterogeneity across sites can be modeled using the Gamma distribution or by a HMM.
- Combining HMMs and phylogenetic trees gives rise to phylo-HMMs, a powerful model for sequence data with many applications.

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