

# AN EXTENSION OF PRATELLI'S INEQUALITY

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ABSTRACT. Let  $X, Y$  be càdlàg martingales and let  $Y^\#$  denote the sharp function of  $Y$ . The paper contains the proof of the estimate

$$\left\| \int_0^\infty |d\langle X, Y \rangle_t| \right\|_1 \leq \sqrt{2} \left\| \langle X \rangle^{1/2} Y^\# \right\|_1$$

for the total variation between  $X$  and  $Y$ . The constant  $\sqrt{2}$  is shown to be the best possible. The proof rests on the construction of an appropriate special function, enjoying certain size and concavity requirements.

## 1. INTRODUCTION

Let  $(\Omega, \mathcal{F}, \mathbb{P})$  be a probability space, equipped with the discrete-time filtration  $(\mathcal{F}_n)_{n \geq 0}$ . Suppose further that  $f = (f_n)_{n \geq 0}$ ,  $g = (g_n)_{n \geq 0}$  are two real-valued and uniformly integrable martingales, with the corresponding pointwise limits denoted by  $f_\infty$  and  $g_\infty$ . Then  $df = (df_n)_{n \geq 0}$ , the difference sequence of  $f$ , is defined by  $df_0 = f_0$  and  $df_n = f_n - f_{n-1}$ . The conditional square function  $s(f)$  of  $f$  is given by the formula  $s(f) = \left( \sum_{n \geq 0} \mathbb{E}(|df_n|^2 | \mathcal{F}_{n-1}) \right)^{1/2}$ , with the convention  $\mathcal{F}_{-1} = \mathcal{F}_0$ . We will also use the truncated (or localized) version  $s_m(f)$ , defined by  $s_m(f) = \left( \sum_{n=0}^m \mathbb{E}(|df_n|^2 | \mathcal{F}_{n-1}) \right)^{1/2}$ . The martingale  $f$  belongs to the space  $h^1$  if and only if its conditional square function is integrable; we define the corresponding  $h^1$ -norm by  $\|f\|_{h^1} := \|s(f)\|_{L^1}$ . It is well-known that  $h^1$  is contained strictly in the usual Hardy space  $H^1$ , defined as the class of martingales  $f$  whose maximal function  $\sup_{n \geq 0} |f_n|$  lies in  $L^1$ . More specifically, there is a finite universal constant  $C$  such that  $\|f\|_{H^1} := \mathbb{E} \sup_{n \geq 0} |f_n| \leq C \|f\|_{h^1}$ , but the reverse bound fails to hold, unless some additional conditions on the filtration are imposed. See Section 2.2 in [22] for details, consult also [8, 13, 21] for more on the subject.

Following Garsia [8], the martingale  $g$  belongs to the class  $BMO_2$ , the space of martingales of bounded mean oscillation, if the quantity

$$\|g\|_{BMO_2} = \sup_{n \geq 0} \left\| \mathbb{E}(|g_\infty - g_{n-1}|^2 | \mathcal{F}_n) \right\|_\infty^{1/2}$$

is finite (we use the convention  $g_{-1} = g_0$ ). This space is dual to  $H^1$  (cf. [8]). (For the analytic versions of the above definition and duality, consult John and Nirenberg [12] and Fefferman [7]). In particular, there is a finite constant  $C$  such that we have

$$|\mathbb{E} f_n g_n| \leq C \|f\|_{H^1} \|g\|_{BMO_2}, \quad n = 0, 1, 2, \dots$$

Pratelli [18] established a version of this estimate, in which  $H^1$  is replaced by the smaller space  $h^1$ . Consider the so-called little space  $bmo_2$ , which consists of all martingales  $g$  for

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which

$$\|g\|_{bmo_2} = \sup_{n \geq 0} \|\mathbb{E}(|g_\infty - g_n|^2 | \mathcal{F}_n)\|_\infty^{1/2} < \infty$$

(note the subtle difference in comparison to  $BMO_2$ : under the expectation, the variable  $g_n$ , instead of  $g_{n-1}$ , is subtracted). Then  $bmo_2$  contains  $BMO_2$  properly (cf. Herz [11]). The aforementioned result of Pratelli asserts that

$$(1.1) \quad \left\| \sum_{n>0} |df_n| |dg_n| \right\|_{L^1} \leq \sqrt{2} \|f\|_{h^1} \|g\|_{bmo_2}.$$

The purpose of this paper is to strengthen this statement. First, we will prove that the constant  $\sqrt{2}$  is the best possible. Second, we will show that the product  $\|f\|_{h^1} \|g\|_{bmo_2}$  can be replaced by the  $L^1$ -norm of a single random variable depending in a bilinear manner on  $f$  and  $g$ . To explain this precisely, let us recall the definition of the sharp maximal function  $g^\#$  of an  $L^2$ -bounded martingale  $g$ . It is given by the formula

$$g^\# = \sup_{n \geq 0} \left( \mathbb{E}(|g_\infty - g_n|^2 | \mathcal{F}_n) \right)^{1/2},$$

and its analytic version appears in various contexts of interpolation theory and, in particular, is of significant importance for the boundedness properties of various classes of operators. For a unified treatment of this topic, we refer the interested reader to Chapter 7 in [10] or Chapter IV in [20].

Our purpose is to establish the following extension of (1.1).

**Theorem 1.1.** *For any uniformly integrable martingale  $f$  and any  $L^2$ -bounded martingale  $g$ , we have*

$$(1.2) \quad \left\| \sum_{n>0} |df_n| |dg_n| \right\|_{L^1} \leq \sqrt{2} \|s(f)g^\#\|_{L^1}.$$

*The constant  $\sqrt{2}$  is the best possible, it is already optimal in (1.1).*

Note that (1.2) does generalize (1.1): indeed, if  $g \in bmo_2$ , then  $g^\# \leq \|g\|_{bmo_2}$  almost surely and hence  $\|s(f)g^\#\|_{L^1} \leq \|f\|_{h^1} \|g\|_{bmo_2}$ .

The above statement should be compared to analogous estimates obtained in [16, 17] which, however, do not involve the sharp function under the  $L^1$ -norm. Furthermore, it should be emphasized that the results in [16, 17] were established under additional assumptions on the regularity of the processes and/or under a different norming of the Hardy space. Here we do not impose any conditions on the filtration. Thus, by a standard discretization argument (see e.g. Remark 11.1.8 and Lemma 11.1.9 in [3]), we can extend the result to a general, continuous-time setting. Let us briefly discuss this possibility. Assume that  $(\Omega, \mathcal{F}, \mathbb{P})$  is a complete probability space, equipped with a right-continuous filtration  $(\mathcal{F}_t)_{t \geq 0}$  such that  $\mathcal{F}_0$  contains all the  $\mathbb{P}$ -null sets of  $\mathcal{F}$ . Suppose that  $X = (X_t)_{t \geq 0}$ ,  $Y = (Y_t)_{t \geq 0}$  are adapted, uniformly integrable martingales, whose trajectories are right-continuous and have limits from the left. For a locally square integrable  $X$ , we denote by  $\langle X \rangle = (\langle X \rangle_t)_{t \geq 0}$  the associated skew bracket, see Chapters VI and VII in Dellacherie and Meyer [5] or Chapter 4 in Métivier [14] for details. Then  $\langle \cdot, \cdot \rangle$  is the quadratic form obtained by polarization and  $\int_0^\infty |d\langle \cdot, \cdot \rangle|$  denotes the associated total

variation. For an  $L^2$ -bounded martingale  $Y$ , its sharp function is given by

$$Y^\# = \sup_{t \geq 0} \left( \mathbb{E}((Y_\infty - Y_t)^2 | \mathcal{F}_t) \right)^{1/2}.$$

Then we have the following version of Theorem 1.1. For related statements with applications to financial mathematics, see e.g. [2, 4].

**Theorem 1.2.** *For all  $X, Y$  as above, we have the sharp estimate*

$$(1.3) \quad \left\| \int_0^\infty |d\langle X, Y \rangle| \right\|_{L^1} \leq \sqrt{2} \left\| \langle X \rangle^{1/2} Y^\# \right\|_{L^1}.$$

*The constant  $\sqrt{2}$  is the best possible. It is already optimal in the estimate*

$$(1.4) \quad \left\| \int_0^\infty |d\langle X, Y \rangle| \right\|_{L^1} \leq \sqrt{2} \left\| \langle X \rangle^{1/2} \right\|_{L^1} \|Y^\#\|_{L^\infty},$$

*even if  $X$  and  $Y$  are assumed to have continuous trajectories.*

We take the advantage to mention here another related result, obtained in [1, Theorem 1.1 (iii)], [6, Lemma 1.6] and [9, Corollary 5.19], under some additional assumptions on the filtration and for continuous martingales:

$$\left\| \int_0^\infty |d\langle X, Y \rangle| \right\|_{L^p} \leq \sqrt{2p} \left\| \langle X \rangle^{1/2} \right\|_{L^p} \|Y^\#\|_{L^\infty}.$$

Furthermore, the order  $O(\sqrt{p})$  as  $p \rightarrow \infty$  in the above estimate is optimal.

Consider the following application of (1.3): if  $1 < p \leq 2$  and  $Y \in L^p$  has mean zero, then we have

$$\begin{aligned} \|Y\|_{L^p} &= \sup_{\|X\|_{L^{p'}} \leq 1} \mathbb{E} X_\infty Y_\infty \leq \sup_{\|X\|_{L^{p'}} \leq 1} \left\| \int_0^\infty |d\langle X, Y \rangle| \right\|_{L^1} \\ &\leq \sup_{\|X\|_{L^{p'}} \leq 1} \sqrt{2} \left\| \langle X \rangle^{1/2} Y^\# \right\|_{L^1} \\ &\leq \sup_{\|X\|_{L^{p'}} \leq 1} \sqrt{2} \left\| \langle X \rangle^{1/2} \right\|_{L^{p'}} \|Y^\#\|_{L^p} \leq \sqrt{\frac{p}{p-1}} \|Y^\#\|_{L^p}. \end{aligned}$$

Here the last inequality follows from the bound  $\left\| \langle X \rangle^{1/2} \right\|_{L^{p'}} \leq \sqrt{p'/2} \|X\|_{L^p}$  (see [21]). Note that the estimate  $\|Y\|_{L^p} \leq c_p \|Y^\#\|_{L^p}$  fails for  $p > 2$ , even for discrete-time martingales. To see this, fix such a  $p$ , take an arbitrary  $\varepsilon > 0$  and consider the martingale  $g = (g_0, g_1, g_1, g_1, \dots)$  with  $g_0 \equiv 0$  and  $\mathbb{P}(g_1 = -\varepsilon) = 1 - \mathbb{P}(g_1 = \varepsilon^{-1}) = (1 + \varepsilon^2)^{-1}$ . Then  $g^\# = \mathbb{E}|g_1|^2 = 1$  almost surely and  $\|g\|_{L^p} \gtrsim \varepsilon^{-1+2/p} \rightarrow \infty$  as  $\varepsilon \rightarrow 0$ .

Our proof of (1.2) will exploit Burkholder's method (sometimes referred to as the Bellman function method): we will deduce the validity of the estimate from the existence of a certain special function enjoying an appropriate concavity. The sharpness of (1.4) will be obtained by the explicit construction of extremal examples.

## 2. PROOF OF THEOREMS 1.1 AND 1.2

**2.1. Proof of (1.2) and (1.3).** A central role in our considerations is played by the special function  $B : (0, \infty) \times \mathbb{R}^3 \rightarrow \mathbb{R}$ , given by the formula

$$B(w, x, y, z) = \frac{x(y^2 - z + 2w^2)}{w}.$$

We will show that this object enjoys the following property.

**Lemma 2.1.** *For any  $w > 0$  and  $x, y, z \in \mathbb{R}$  such that  $0 \leq z - y^2 \leq w^2$  and any  $d, e, f \in \mathbb{R}$  we have the inequality*

$$(2.1) \quad \begin{aligned} & B(w \vee (z + f - (y + e)^2)_+^{1/2}, (x^2 + d^2)^{1/2}, y + e, z + f) \\ & \geq B(w, x, y, z) + (x^2 + d^2)^{1/2} \frac{2ye - f}{w} + \frac{d^2 w}{2(x^2 + d^2)^{1/2}} + (x^2 + d^2)^{1/2} \frac{e^2}{w}. \end{aligned}$$

*Proof.* We consider two cases. If  $z + f - (y + e)^2 < w^2$ , then the estimate is equivalent to

$$\left( (x^2 + d^2)^{1/2} - x \right) \frac{y^2 - z + 2w^2}{w} \geq \frac{d^2 w}{2(x^2 + d^2)^{1/2}},$$

which is obvious: we have  $y^2 - z + 2w^2 \geq w^2$  and  $(x^2 + d^2)^{1/2} - x \geq d^2 / (2(x^2 + d^2)^{1/2})$ . If the reverse inequality  $z + f - (y + e)^2 \geq w^2$  holds, then we rewrite (2.1) in the form

$$\begin{aligned} & (x^2 + d^2)^{1/2} \left[ (z + f - (y + e)^2)^{1/2} + \frac{z + f - (y + e)^2 - 2w^2}{w} \right] \\ & \geq (x - (x^2 + d^2)^{1/2}) \frac{y^2 - z + 2w^2}{w} + \frac{d^2 w}{2(x^2 + d^2)^{1/2}}. \end{aligned}$$

Observe that the left-hand side decreases as  $e$  increases: therefore, it is enough to show the claim for  $z + f - (y + e)^2 = w^2$ , which follows from the previous case and continuity. Let us observe that by Schwarz' inequality, (2.1) gives the slightly weaker bound

$$(2.2) \quad \begin{aligned} & B(w \vee (z + f - (y + e)^2)_+^{1/2}, (x^2 + d^2)^{1/2}, y + e, z + f) \\ & \geq B(w, x, y, z) + (x^2 + d^2)^{1/2} \frac{2ye - f}{w} + \sqrt{2}|d||e|, \end{aligned}$$

which will be more suitable for our purposes below.  $\square$

*Proof of (1.2) and (1.3).* It is enough to establish (1.2), the second estimate follows by standard discretization. Fix  $f, g$  as in the statement and introduce the auxiliary sequences  $h_n = \mathbb{E}(g^2 | \mathcal{F}_n)$  and  $w_n = \max_{0 \leq k \leq n} (h_k - g_k^2)^{1/2}$  for  $n = 0, 1, 2, \dots$  (note that  $h_k \geq g_k^2$  by Schwarz' inequality, so  $w_n$  is well-defined). We may assume that  $w_0 > 0$  almost surely (and hence also  $\mathbb{P}(w_n > 0) = 1$  for all  $n$ ), by simple perturbation argument. The key observation is that for any  $n \geq 1$  we have

$$(2.3) \quad \mathbb{E}B(w_n, s_n(f), g_n, h_n) \geq \mathbb{E}B(w_{n-1}, s_{n-1}(f), g_{n-1}, h_{n-1}) + \sqrt{2}\mathbb{E}|df_n||dg_n|.$$

To see this, we apply (2.2) with  $w = w_{n-1}$ ,  $x = s_{n-1}(f)$ ,  $y = g_{n-1}$ ,  $z = h_{n-1}$  and  $d = (\mathbb{E}(|df_n|^2 | \mathcal{F}_{n-1}))^{1/2}$ ,  $e = dg_n$ ,  $f = dh_n$ . The required condition  $0 \leq z - y^2 \leq w$  follows from Schwarz' inequality and the definition of the sharp function. Integrating both sides gives (2.3), since  $g, h$  are martingales. Therefore, by induction, we obtain

$$\begin{aligned} \sqrt{2} \left\| \sum_{n=1}^N |df_n||dg_n| \right\|_{L^1} & \leq \sqrt{2}\mathbb{E} \sum_{n=1}^N |df_n||dg_n| + \mathbb{E}B(w_0, s_0(f), g_0, h_0) \\ & \leq \mathbb{E}B(w_N, s_N(f), g_N, h_N). \end{aligned}$$

Since  $h_N \geq g_N^2$ , we have  $\mathbb{E}B(w_N, s_N(f), g_N, h_N) \leq 2\mathbb{E}s_N(f)w_N \leq 2\mathbb{E}s(f)g^\#$  and the estimate follows.  $\square$

**2.2. Sharpness.** We turn our attention to the optimality of the constant  $\sqrt{2}$  in (1.1), (1.2), (1.3) and (1.4). Of course, it is enough to focus on the last estimate. Let  $\delta, \lambda \in (0, 1)$  be fixed parameters. We start with a well-known property of Brownian motion (see e.g. [19], p. 73), which can be proved by a simple martingale argument.

**Lemma 2.2.** *Let  $Y$  be a standard Brownian motion started at zero. Then the stopping time  $\sigma = \inf\{t > 0 : |Y_t| = \delta\}$  satisfies  $\mathbb{E} \exp(\lambda \sigma) = (\cos(\sqrt{2\lambda}\delta))^{-1}$ .*

We are ready for the construction of extremal processes in (1.4). Let  $W = (W_t)_{t \geq 0}$  be a standard Brownian motion and introduce the stopping time  $\tau_0 \equiv 0$ . To describe the appropriate *BMO* martingale, we will construct first an appropriate two-dimensional martingale, taking values in a parabolic domain  $\{(y, z) : y^2 \leq z \leq y^2 + 1\}$ , and then restrict ourselves to the first coordinate of this process (this representation goes back to the works [15, 16]). So, consider the two-dimensional martingale  $(Y, Z)$ , starting from  $(0, 1)$ , whose evolution is governed by the inductive use of the following two rules.

1° Suppose that we have successfully defined  $\tau_{2n}$ . If  $(Y_{\tau_{2n}}, Z_{\tau_{2n}})$  belongs to the parabola  $z = y^2$ , then the process  $(Y, Z)$  does not move and we set  $\tau_{2n+1} := \tau_{2n}$ . On the other hand, suppose that  $(Y_{\tau_{2n}}, Z_{\tau_{2n}})$  lies on the parabola  $z = y^2 + 1$  and denote  $y_0 = Y_{\tau_{2n}}$ ,  $z_0 = Z_{\tau_{2n}}$ . Then for  $t > \tau_{2n}$ , the pair  $(Y_t, Z_t)$  evolves along the line segment passing through  $(y_0, z_0)$  and tangent to the parabola  $z = y^2 + 1$ , until  $Y$  gets to  $y_0 - \delta$  or  $y_0 + \delta$ . More precisely, we let  $Y_t = y_0 + (W_t - W_{\tau_{2n}})$  and  $Z_t = z_0 + 2y_0(Y_t - y_0)$  for  $t \in (\tau_{2n}, \tau_{2n+1})$ , where  $\tau_{2n+1} = \inf\{t > \tau_{2n} : W_t = W_{\tau_{2n}} \pm \delta\}$ . For the further use, let us record that

$$(2.4) \quad Z_{\tau_{2n+1}} - Y_{\tau_{2n+1}}^2 = z_0 - y_0^2 - (Y_{\tau_{2n+1}} - y_0)^2 = 1 - \delta^2$$

and go to 2°.

2° Suppose that we have successfully defined  $\tau_{2n+1}$ . If  $(Y_{\tau_{2n+1}}, Z_{\tau_{2n+1}})$  belongs to the parabola  $z = y^2$ , then the process  $(Y, Z)$  does not move and we set  $\tau_{2n+2} := \tau_{2n+1}$ . Otherwise, the process  $(Y, Z)$  evolves vertically until it gets to one of the parabolas  $z = y^2$  or  $z = y^2 + 1$ . Formally, for  $t \in (\tau_{2n+1}, \tau_{2n+2}]$  we set  $Y_t = Y_{\tau_{2n+1}}$  and  $Z_t = Z_{\tau_{2n+1}} + (W_t - W_{\tau_{2n+1}})$ , where  $\tau_{2n+2} = \inf\{t > \tau_{2n+1} : W_t - W_{\tau_{2n+1}} \in \{\delta^2 - 1, \delta^2\}\}$ . The identity (2.4) implies that the probability of reaching the upper parabola is equal to  $1 - \delta^2$ . Go to 1°.

Let us gather some information about the pair  $(Y, Z)$  which follows directly from the above construction. First, as we have announced before,  $(Y, Z)$  is a martingale taking values in the parabolic domain  $\{(y, z) : y^2 \leq z \leq y^2 + 1\}$ , in particular,  $Y$  is  $L^2$  bounded: for any  $t \geq 0$  we have  $\mathbb{E}Y_t^2 \leq \mathbb{E}Z_t = \mathbb{E}Z_0 = 1$ . Furthermore, by the final sentence in the description of 2°,

$$\mathbb{P}(Z_{\tau_{2n+2}} > Y_{\tau_{2n+2}}^2 | Z_{\tau_{2n}} > Y_{\tau_{2n}}^2) = 1 - \delta^2.$$

This implies that the pair terminates at the lower parabola  $z = y^2$ , i.e., we have  $Z_\infty = Y_\infty^2$  almost surely: indeed, by the above identity, we see that  $\mathbb{P}(Z_{\tau_{2n}} > Y_{\tau_{2n}}^2) = (1 - \delta^2)^n \rightarrow 0$ . These observations imply that  $Y$  is a *BMO* martingale: we have  $Y^\# \leq 1$  almost surely. To show this, note that for an arbitrary time  $t \geq 0$  we have

$$\mathbb{E}((Y_\infty - Y_t)^2 | \mathcal{F}_t) = \mathbb{E}(Y_\infty^2 | \mathcal{F}_t) - Y_t^2 = \mathbb{E}(Z_\infty | \mathcal{F}_t) - Y_t^2 = Z_t - Y_t^2 \leq 1$$

almost surely. Next, observe that  $Y$  behaves like a Brownian motion on the intervals  $[\tau_{2n}, \tau_{2n+1}]$  and is constant on the intervals  $[\tau_{2n+1}, \tau_{2n+2}]$ ,  $n = 0, 1, 2, \dots$ . Consequently,

for  $t \in [\tau_{2n}, \tau_{2n+1}]$  we have

$$\langle Y \rangle_t = t - \tau_{2n} + \sum_{k=0}^{n-1} (\tau_{2k+1} - \tau_{2k}).$$

Let  $\tau = \lim_{n \rightarrow \infty} \tau_n$  be the lifetime of  $(Y, Z)$  and define  $N = \inf\{n : \tau_{2n+2} = \tau\}$ . Moreover, put  $\sigma_n = \tau_{2n+1} - \tau_{2n}$  for  $n = 0, 1, 2, \dots$ . Then  $N$  follows the geometric law:  $\mathbb{P}(N \geq n) = (1 - \delta^2)^n$  for  $n = 0, 1, 2, \dots$ . Furthermore, by standard properties of Brownian motion and Lemma 2.2, we have

$$\mathbb{E} \exp \left( \lambda \sum_{n=0}^N \sigma_n \right) = \sum_{n=0}^{\infty} (\cos(\sqrt{2\lambda}\delta))^{-n-1} (1 - \delta^2)^n \delta^2.$$

This can be checked by conditioning with respect to  $N$ : indeed, conditionally on the set  $\{N = n\}$ , the random variables  $\sigma_0, \sigma_1, \sigma_2, \dots, \sigma_n$  are independent copies of the variable  $\sigma$  appearing in Lemma 2.2. Note that the series on the right can be made arbitrarily large, by taking first  $\lambda$  sufficiently close to 1 and then  $\delta$  close to zero: indeed,  $\cos x = 1 - x^2/2 + o(x^2)$  as  $x \rightarrow 0$ .

We proceed to the definition of the process  $X$ , which is given by the stochastic integral

$$X_t = \int_{0+}^t e^{\lambda \langle Y \rangle_s} dY_s, \quad t \geq 0.$$

By the above discussion, we have

$$\langle X \rangle_{\infty} = \sum_{n=0}^{\infty} \int_{\tau_{2n}}^{\tau_{2n+1}} e^{2\lambda \langle Y \rangle_s} ds \leq \frac{1}{2\lambda} \exp \left( 2\lambda \sum_{n=0}^N \sigma_n \right)$$

and hence

$$\mathbb{E} \langle X \rangle_{\infty}^{1/2} \leq \frac{1}{\sqrt{2\lambda}} \mathbb{E} \exp \left( \lambda \sum_{n=0}^N \sigma_n \right).$$

A similar calculation shows that

$$\mathbb{E} \int_0^{\infty} |d\langle X, Y \rangle_t| = \mathbb{E} X_{\infty} Y_{\infty} = \mathbb{E} \sum_{n=0}^{\infty} \int_{\tau_{2n}}^{\tau_{2n+1}} e^{\lambda \langle Y \rangle_s} ds = \frac{\mathbb{E} \left( \exp \left( \lambda \sum_{n=0}^N \sigma_n \right) - 1 \right)}{\lambda}.$$

However, as we have discussed above, for appropriately chosen  $\lambda$  and  $\delta$ , the expectation  $\mathbb{E} \exp \left( \lambda \sum_{n=0}^N \sigma_n \right)$  can be made arbitrarily large. This immediately implies that the ratio  $\mathbb{E} \int_0^{\infty} |d\langle X, Y \rangle_t| / \mathbb{E} \langle X \rangle_{\infty}^{1/2}$  can be made as close to  $\sqrt{2}$  as we wish. This is precisely the desired sharpness of (1.4).

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