

# Smooth functionals of covariance operators: minimax estimation error bounds and effective rank

Vladimir Koltchinskii

School of Mathematics  
Georgia Institute of Technology

Bedlewo, 2023

# Covariance and sample covariance operators

- $X \sim N(0, \Sigma)$  a centered Gaussian r.v. in a separable Banach space  $E$  with the dual space  $E^*$
- $\Sigma : E^* \mapsto E$  is the covariance operator of  $X$  :

$$\Sigma u := \mathbb{E} \langle X, u \rangle X, u \in E^*$$

- $X_1, \dots, X_n$  i.i.d. copies of  $X$
- $\hat{\Sigma} = \hat{\Sigma}_n : E^* \mapsto E$  is the sample covariance operator based on  $X_1, \dots, X_n$

$$\hat{\Sigma} u := n^{-1} \sum_{j=1}^n \langle X_j, u \rangle X_j, u \in E^*$$

# Covariance operators and effective rank

- The **effective rank** of  $\Sigma$  : for  $X \sim N(0, \Sigma)$ ,

$$\begin{aligned} \mathbf{r}(\Sigma) &:= \frac{\mathbb{E}\|X\|^2}{\|\Sigma\|} \\ &= \frac{\mathbb{E} \sup_{\|u\|, \|v\| \leq 1} \langle X, u \rangle \langle X, v \rangle}{\sup_{\|u\|, \|v\| \leq 1} \mathbb{E} \langle X, u \rangle \langle X, v \rangle} \geq 1 \end{aligned}$$

- $\mathbf{r}(\lambda\Sigma) = \mathbf{r}(\Sigma)$ ,  $\lambda > 0$
- $\mathbf{r}(\Sigma) \leq \text{rank}(\Sigma) \leq \dim(E)$
- If  $E = \mathbb{H}$  is a Hilbert space, then  $\mathbf{r}(\Sigma) = \frac{\text{tr}(\Sigma)}{\|\Sigma\|}$
- If  $E = \mathbb{R}^d$  with the Euclidean norm and  $\sigma(\Sigma) \subset [a^{-1}, a]$  for some  $a \geq 1$  ("almost isotropic covariance"), then  $\mathbf{r}(\Sigma) \asymp d$ .

# Gaussian Version of Dvoretzky's Theorem and Effective Rank

## Theorem (Pisier (1986, 1989))

*For all  $\varepsilon \in (0, 1)$ , there exists  $\eta(\varepsilon) > 0$  with the following property. If, in a Banach space  $E$ , there exists  $X \sim N(0, \Sigma)$  with  $\mathbf{r}(\Sigma) = r$ , then  $E$  contains a subspace  $F$  of dimension  $\sim \eta(\varepsilon)r$  which is  $(1 \pm \varepsilon)$ -isomorphic to  $\ell_2^n$ .*

**Pisier (1989)** called  $\mathbf{r}(\Sigma)$  “the dimension of  $X$ ”, or “the concentration dimension” of  $X$ .

# Bounding the sample covariance via effective rank

## Theorem (K& Lounici (2014))

Let  $X \sim N(0, \Sigma)$  in  $E$  and let  $X_1, \dots, X_n$  be i.i.d. copies of  $X$ . Then

$$\mathbb{E}\|\hat{\Sigma} - \Sigma\| \asymp \|\Sigma\| \left( \sqrt{\frac{\mathbf{r}(\Sigma)}{n}} \vee \frac{\mathbf{r}(\Sigma)}{n} \right),$$

and, for all  $t \geq 1$  with probability  $\geq 1 - e^{-t}$ ,

$$\left| \|\hat{\Sigma} - \Sigma\| - \mathbb{E}\|\hat{\Sigma} - \Sigma\| \right| \lesssim \|\Sigma\| \left( \sqrt{\frac{\mathbf{r}(\Sigma)}{n}} \vee 1 \right) \sqrt{\frac{t}{n}} \vee \|\Sigma\| \frac{t}{n}.$$

- **Earlier results:** in the case  $E = \mathbb{R}^d$  with the Euclidean norm, bounds with  $\log d$ -factors based on non-commutative Khintchine inequalities by [Lust-Piquard and Pisier](#) ([Vershynin](#), around 2011-2012).
- **Alternative proof** of concentration bound: [Adamczak \(2014\)](#)

# Functionals of covariance

- $L(E^*, E)$  the space of bounded symmetric linear operators from  $E^*$  into  $E$  equipped with the operator norm
- $f : L(E^*, E) \mapsto \mathbb{R}$  a functional
- $f(\Sigma)$  to be estimated based on i.i.d.  $X_1, \dots, X_n \sim N(0, \Sigma)$
- For  $a > 0$  and  $r \geq 1$ ,

$$\mathcal{S}(a, r) := \{\Sigma : \|\Sigma\| \leq a, \mathbf{r}(\Sigma) \leq r\}$$

- Find the size of

$$\sup_{\|f\|_{C^s} \leq 1} \inf_{T_n} \sup_{\Sigma \in \mathcal{S}(a, r)} \left\| T_n(X_1, \dots, X_n) - f(\Sigma) \right\|_{L_2(\mathbb{P}_\Sigma)}.$$

- Is it possible to construct an estimator with  $\sqrt{n}$ -rate?  
Asymptotically normal estimator? Asymptotically efficient estimator?

# Hölder smoothness

- $F$  is a Banach space (say,  $F = L(E^*, E)$ ),  $\Theta \subset F$  an open subset
- $g : \Theta \mapsto F_1$ ,  $F_1$  is a Banach space

$$\|g\|_{L_\infty(\Theta)} := \sup_{x \in \Theta} \|g(x)\|, \quad \|g\|_{\text{Lip}(\Theta)} := \sup_{x, x' \in \Theta, x \neq x'} \frac{\|g(x) - g(x')\|}{\|x - x'\|}$$

$$\|g\|_{\text{Lip}_\rho(\Theta)} := \sup_{x, x' \in \Theta, x \neq x'} \frac{\|g(x) - g(x')\|}{\|x - x'\|^\rho}, \quad \rho \in (0, 1].$$

- $g : \Theta \mapsto \mathbb{R}$   $k$  times Fréchet differentiable for some  $k \geq 0$
- For  $s = k + \rho$  with  $\rho \in (0, 1]$ , define

$$\|g\|_{C^s(\Theta)} := \max\left(\|g\|_{L_\infty}, \max_{0 \leq j \leq k-1} \|g^{(j)}\|_{\text{Lip}}, \|g^{(k)}\|_{\text{Lip}_\rho}\right).$$

- **Note:** norms of the derivatives are defined as the operator norms (of multilinear forms).
- $C^s(\Theta) := \{g : \Theta \mapsto \mathbb{R} : \|g\|_{C^s(\Theta)} < \infty\}$

# Example of Hölder smoothness: functions of operators

- $E = E^* = \mathbb{H}$ ,  $\mathbb{H}$  is a Hilbert space
- $\mathcal{L}_{sa}(\mathbb{H})$  is the space of self-adjoint operators in  $\mathbb{H}$  equipped with the operator norm
- $g : \mathbb{R} \mapsto \mathbb{R}$  induces  $\mathcal{L}_{sa}(\mathbb{H}) \ni A \mapsto g(A) \in \mathcal{L}_{sa}(\mathbb{H})$
- for a compact operator  $A$ ,

$$A = \sum_{\lambda \in \sigma(A)} \lambda P_\lambda \implies g(A) = \sum_{\lambda \in \sigma(A)} g(\lambda) P_\lambda$$

- K (2017):

$$\|g(\cdot)\|_{C^s(\mathcal{L}_{sa}(\mathbb{H}))} \lesssim_s \|g\|_{B_{\infty,1}^s(\mathbb{R})}$$

- Then, for  $f(A) := \text{tr}(g(A)B)$ ,  $A \in \mathcal{L}_{sa}(\mathbb{H})$  and  $\|B\|_1 < \infty$

$$\|f\|_{C^s(\mathcal{L}_{sa}(\mathbb{H}))} \lesssim_s \|g\|_{B_{\infty,1}^s(\mathbb{R})} \|B\|_1$$



# Example of Hölder smoothness: functions of operators

- $\Sigma_0 : \mathbb{H} \mapsto \mathbb{H}$  a covariance operator with eigenvalues  $\|\Sigma_0\| = \lambda_1 = \dots = \lambda_l > \lambda_{l+1} \geq \dots$
- $g_l := \lambda_l - \lambda_{l+1}$  the spectral gap
- $U = B(\Sigma_0; \delta) := \{A : \|A - \Sigma_0\| < \delta\}$ ,  $\delta < g_l/8$ .
- $P(A)$  the orthogonal projection onto the linear span of eigenvectors corresponding to the top  $l$  eigenvalues of  $A$
- The function  $U \ni A \mapsto P(A)$  is  $C^\infty$  and

$$\|P^{(k)}\|_{L_\infty(U)} \lesssim_k g_l^{-k}, k \geq 0$$

- For  $f(A) := \langle P(A), B \rangle$ ,

$$\|f^{(k)}\|_{L_\infty(U)} \lesssim_k \|B\|_1 g_l^{-k}, k \geq 0$$

# A classical problem

- $X_1, \dots, X_n$  i.i.d.  $\sim P_\theta, \theta \in \Theta, \Theta \subset \mathbb{R}^d$  an open set,  $d \geq 1, n \rightarrow \infty$
- $\{P_\theta : \theta \in \Theta\}$  a regular statistical model with non-singular Fisher information  $I(\theta)$
- $f : \Theta \mapsto \mathbb{R}$  a continuously differentiable function
- $\hat{\theta}_n$  maximum likelihood estimator (MLE) based on  $X_1, \dots, X_n$
- Then, uniformly in  $\Theta$ ,

$$n\mathbb{E}_\theta(f(\hat{\theta}_n) - f(\theta))^2 \rightarrow \sigma_f^2(\theta) := \langle I(\theta)^{-1} f'(\theta), f'(\theta) \rangle$$

and

$$\sqrt{n}(f(\hat{\theta}_n) - f(\theta)) \xrightarrow{d} N(0; \sigma_f^2(\theta)) \text{ as } n \rightarrow \infty$$

- Moreover,  $f(\hat{\theta}_n)$  is a locally asymptotically minimax estimator of  $f(\theta)$  in the sense of Hájek and Le Cam:

$$\lim_{c \rightarrow \infty} \liminf_{n \rightarrow \infty} \inf_{T_n} \sup_{\|\theta - \theta_0\| \leq cn^{-1/2}} n\mathbb{E}_\theta(T_n(X_1, \dots, X_n) - f(\theta))^2 \geq \sigma_f^2(\theta_0).$$

# Previous results: high-dimensional and non-parametric models

- Ibragimov, Nemirovski and Khasminskii (1987), Nemirovski (1990, 2000)
- estimation of  $f(\theta)$  for functionals  $f$  of smoothness  $s$ ,  $\theta$  being the parameter of infinite-dimensional Gaussian white noise model:

$$dX(t) = \theta(t)dt + n^{-1/2}dw(t), t \in [0, 1], \theta \in \Theta \subset L_2([0, 1])$$

- proved the existence of a “smoothness threshold” such that the efficient estimation with  $\sqrt{n}$ -rate is possible when  $s$  is above the threshold and is impossible (for some functionals) otherwise;
- the smoothness threshold depends on the rate of decay of Kolmogorov widths of the parameter set.

# Previous results: high-dimensional and non-parametric models

- $E = \mathbb{R}^d$  equipped with the Euclidean norm

$$\tilde{\mathcal{S}}(a; d) := \{\Sigma : \sigma(\Sigma) \subset [1/a, a]\}, a \geq 1$$

- $\mathbf{r}(\Sigma) \asymp d, \Sigma \in \tilde{\mathcal{S}}(a; d)$
- [K \(2017, 2018\)](#) studied estimation of functionals  $\langle g(\Sigma), B \rangle$  and determined the degree of smoothness of function  $g : \mathbb{R} \mapsto \mathbb{R}$  needed for asymptotically efficient estimation
- [K& Zhilova \(2019\)](#) developed estimators  $\hat{T}_f(X_1, \dots, X_n)$  for general Hölder smooth functionals  $f$  such that

$$\sup_{\|f\|_{C^s} \leq 1} \sup_{\Sigma \in \tilde{\mathcal{S}}(a, d)} \left\| \hat{T}_f - f(\Sigma) \right\|_{L_2(\mathbb{P}_\Sigma)} \lesssim_a \left( \frac{1}{\sqrt{n}} \vee \left( \sqrt{\frac{d}{n}} \right)^s \right) \wedge 1$$

- If  $d \leq n^\alpha$  for some  $\alpha \in (0, 1)$  and  $s \geq \frac{1}{1-\alpha}$ , then the  $L_2$ -error rate is  $O(n^{-1/2})$  and, for  $s > \frac{1}{1-\alpha}$ ,  $\hat{T}_f$  is asymptotically efficient

# Previous results: high-dimensional and non-parametric models

The error rate

$$\left( \frac{1}{\sqrt{n}} \vee \left( \sqrt{\frac{d}{n}} \right)^s \right) \wedge 1$$

is typical in the problems of this type:

- finite-dimensional and infinite dimensional random shift models with Gaussian and Poincaré noise: [K& Zhilova \(2018\)](#), [K& Zhilova \(2020\)](#)
- general high-dimensional parametric models with an estimator admitting normal approximation: [K \(2020\)](#)
- high-dimensional log-concave location models: [K and Wahl \(2021\)](#)

The estimation method in these papers was based on higher order iterative bias reduction (“iterated bootstrap”, [Hall and Martin \(1988\)](#))

# Loss functions and Orlicz norm

- $\psi : \mathbb{R}_+ \mapsto \mathbb{R}_+$  be a convex nondecreasing loss function such that  $\psi(0) = 0$
- $\|\eta\|_\psi := \|\eta\|_{L_\psi(\mathbb{P})} := \inf \left\{ c > 0 : \mathbb{E} \psi \left( \frac{|\eta|}{c} \right) \leq 1 \right\}$
- For  $\psi(u) := u^p, p \geq 1$ ,  $\|\eta\|_\psi = \|\eta\|_{L_p}$
- Other choices:  $\psi(u) = \psi_1(u) := e^u - 1$  (subexponential loss) and  $\psi(u) = \psi_2(u) = e^{u^2} - 1$  (subgaussian loss). More generally,  $\psi_\alpha(u) := e^{u^\alpha} - 1, \alpha \geq 1$ .
- $\|\eta\|_{\psi_\alpha} \asymp \sup_{p \geq 1} p^{-1/\alpha} \|\eta\|_{L_p}, \alpha > 0$
- Given two loss functions  $\psi$  and  $\varphi$ , we write  $\psi \preceq \varphi$  ( $\psi$  is dominated by  $\varphi$ ) iff

$$\psi(u) \leq c_1 \varphi(c_2 u), u \geq 0$$

for some constants  $c_1, c_2 > 0$ .

# Upper bound on the $L_\psi$ -risk (K (2022))

For  $a > 0$ ,  $U \subset L(E^*, E)$  and  $s = k + \rho$ ,  $k \geq 0$ ,  $\rho \in (0, 1]$ , define the **weighted Hölder norm** of  $f$  as

$$\|f\|_{C^{s,a}(U)} := \max_{0 \leq j \leq k} a^j \|f^{(j)}\|_{L_\infty(U)} \vee a^s \|f^{(k)}\|_{\text{Lip}_\rho(U)}.$$

## Theorem

Let  $a > 0$ ,  $r \geq 1$  and  $\Sigma_0 \in \mathcal{S}(a, r)$ . Suppose, for a sufficiently large  $C > 0$ ,  $Ca\sqrt{\frac{r}{n}} < \delta$ . Let  $U := B(\Sigma_0, 2\delta) := \{\Sigma : \|\Sigma - \Sigma_0\| < 2\delta\}$  and let  $s = k + \rho$ ,  $k \geq 0$ ,  $\rho \in (0, 1]$ . Then, there exists an estimator  $T_f(X_1, \dots, X_n)$  of  $f(\Sigma)$  such that, for all  $\psi \preceq \psi_1$

$$\begin{aligned} & \sup_{\|f\|_{C^{s,a}(U)} \leq 1} \sup_{\Sigma \in \mathcal{S}(a,r), \|\Sigma - \Sigma_0\| < \delta} \left\| T_f(X_1, \dots, X_n) - f(\Sigma) \right\|_{L_\psi(\mathbb{P}_\Sigma)} \\ & \lesssim_s \frac{1}{\sqrt{n}} \vee \left( \sqrt{\frac{r}{n}} \right)^s. \end{aligned}$$

# Minimax lower bound (K (2022))

- $E = \mathbb{H}$  is a separable Hilbert space.
- $\Sigma_0$  is a spiked covariance of rank  $d \iff \sigma(\Sigma_0) = \{\lambda, \mu, 0\}$  with  $\lambda > \mu > 0$ ,  $\lambda$  of multiplicity 1,  $\mu$  of multiplicity  $d - 1$ .

## Theorem

Let  $\Sigma_0 \in \mathcal{S}(a, r)$  be a spiked covariance of rank  $[r]$  with  $\lambda := \gamma_1 \mathbf{a}$ ,  $\mu := \gamma_2 \mathbf{a}$ ,  $0 < \gamma_2 \leq \gamma_1$ . Denote  $\kappa := \gamma_2 \wedge (\gamma_1 - \gamma_2) \wedge (1 - \gamma_1)$  and suppose that  $c_1 \gamma_1 \mathbf{a} \sqrt{\frac{r}{n}} < \delta \leq c_2 \kappa \mathbf{a} \wedge 1$  for a sufficiently large  $c_1$  and sufficiently small  $c_2$ . Let  $U := B(\Sigma_0, 2\delta)$  and let  $s > 0$ . Then

$$\sup_{\|f\|_{\mathcal{C}^s, \mathbf{a}(U)} \leq 1} \inf_T \sup_{\Sigma \in \mathcal{S}(a, r), \|\Sigma - \Sigma_0\| < \delta} \left\| T(X_1, \dots, X_n) - f(\Sigma) \right\|_{L_2(\mathbb{P}_\Sigma)} \\ \asymp_{\mathbf{s}, \gamma_1, \gamma_2} \frac{1}{\sqrt{n}} \vee \left( \sqrt{\frac{r}{n}} \right)^s.$$



# The lower bound: a sketch of the proof

- $\Sigma_0 = \lambda(u \otimes u) + \mu(P_L - u \otimes u)$   
 $L \subset \mathbb{H}$ ,  $\dim(L) = d := [r]$ ,  $u \in L$ ,  $\|u\| = 1$   
( $L$  could be identified with  $\mathbb{R}^d$ )
- $\theta(\Sigma)$ ,  $\Sigma \in U$  the eigenvector corresponding to the top eigenvalue of  $\Sigma$ ,  $\langle \theta(\Sigma), u \rangle \geq 0$ ,  $U \ni \Sigma \mapsto \theta(\Sigma)$  is  $C^\infty$
- We will construct certain **least favorable functionals**

$$f_k(\Sigma) := h_k(\theta(\Sigma)), \Sigma \in U, k = 1, \dots, d,$$

where  $h_k : \mathbb{H} \mapsto \mathbb{R}$ ,  $\|h_k\|_{C^s} \lesssim 1$

- For these functionals,  $\|f_k\|_{C^{s,a}(U)} \leq 1$  and we will show that

$$\begin{aligned} & \max_{1 \leq k \leq d} \inf_{T_k} \sup_{\Sigma \in S(a,r), \|\Sigma - \Sigma_0\| < \delta} \left\| T_k(X_1, \dots, X_n) - f_k(\Sigma) \right\|_{L_2(\mathbb{P}_\Sigma)} \\ & \gtrsim_{S, \gamma_1, \gamma_2} \left( \sqrt{\frac{r}{n}} \right)^s. \end{aligned}$$

# Well separated subsets

- $\exists B \subset \{-1, 1\}^d : \text{card}(B) \geq \frac{e^{d/8}}{2}, |\langle u, \omega \rangle| < 2$  and

$$h(\omega, \omega') := \sum_{j=1}^d I(\omega_j \neq \omega'_j) \geq \frac{d}{4}, \omega, \omega' \in B, \omega \neq \omega'.$$

- Let  $\varepsilon \asymp \sqrt{\frac{d}{n}}$
- $\Theta_\varepsilon = \{\theta_\omega : \omega \in B\}, \theta_\omega := \frac{t_\omega}{\|t_\omega\|}, t_\omega := \varepsilon \frac{\omega}{\sqrt{d}} + \sqrt{1 - \varepsilon^2} u, \omega \in B$

$$\frac{\varepsilon}{2\sqrt{d}} \sqrt{h(\omega, \omega')} \leq \|\theta_\omega - \theta_{\omega'}\| \leq \frac{8\varepsilon}{\sqrt{d}} \sqrt{h(\omega, \omega')}, \omega, \omega' \in B,$$

implying that

$$\frac{\varepsilon}{4} \leq \|\theta_\omega - \theta_{\omega'}\| \leq 8\varepsilon, \omega, \omega' \in B, \omega \neq \omega'$$

# Well separated subsets

- $\Sigma_\theta := \lambda(\theta \otimes \theta) + \mu(P_L - \theta \otimes \theta)$ ,  $\theta \in L$ ,  $\|\theta\| = 1$
- Then

$$\|\Sigma_{\theta_\omega} - \Sigma_0\| < \delta, \omega \in B$$

- Moreover, for all  $\omega, \omega' \in B$ ,

$$K(N(0, \Sigma_{\theta_\omega})^{\otimes n} \| N(0, \Sigma_{\theta_{\omega'}})^{\otimes n}) \lesssim n\varepsilon^2 \lesssim \log \text{card}(B),$$

implying that for  $X_1, \dots, X_n$  i.i.d.  $\sim N(0, \Sigma_{\theta_\omega})$ ,  $\omega \in B$

$$\inf_{\hat{\theta}} \max_{\omega \in B} \mathbb{E}_{\Sigma_{\theta_\omega}} \|\hat{\theta} - \theta_\omega\|^2 \gtrsim \varepsilon^2.$$

# Nemirovski's bump functionals

- $\varphi : \mathbb{R} \mapsto [0, 1]$ ,  $\varphi$  is  $C^\infty$ ,  $\text{supp}(\varphi) \in [-1, 1]$ ,  $\varphi(0) > 0$
- $\phi(\mathbf{u}) := \varphi(\|\mathbf{u}\|^2)$ ,  $\mathbf{u} \in \mathbb{H}$
- 

$$h_k(\theta) := \sum_{\omega \in B} \omega_k \varepsilon^s \phi\left(\frac{\theta - \theta_\omega}{c\varepsilon}\right), k = 1, \dots, d,$$

$c > 0$  small enough

- Note that “bump functions”  $\phi\left(\frac{\theta - \theta_\omega}{c\varepsilon}\right)$ ,  $\omega \in B$  have disjoint supports and  $\|h_k\|_{C^s} \lesssim 1$ .
- $h_k(\theta_\omega) = \omega_k \varepsilon^s \varphi(0)$ ,  $k = 1, \dots, d$ ,  $\omega \in B$
- The values  $h_k(\theta_\omega)$ ,  $k = 1, \dots, d$  provide a “coding” for  $\theta_\omega$

# Back to least favorable functionals

- $f_k(\Sigma) := h_k(\theta(\Sigma)), \Sigma \in U, k = 1, \dots, d$
- $f_k(\Sigma_{\theta_\omega}) = \omega_k \varepsilon^s \varphi(0), k = 1, \dots, d, \omega \in B$
- Define

$$\tau^2(\omega, \omega') := \frac{1}{d} \sum_{k=1}^d (f_k(\Sigma_{\theta_\omega}) - f_k(\Sigma_{\theta_{\omega'}}))^2, \omega, \omega' \in B.$$

- Then

$$\tau^2(\omega, \omega') \asymp \varepsilon^{2s} \frac{h(\omega, \omega')}{d}, \omega, \omega' \in B,$$

implying that

$$\tau^2(\omega, \omega') \asymp \varepsilon^{2(s-1)} \|\theta_\omega - \theta_{\omega'}\|^2, \omega, \omega' \in B.$$

# Back to lower bounds

- If there exist estimators  $\hat{T}_k$  of  $f_k(\Sigma_{\theta_\omega})$ ,  $k = 1, \dots, d$  based on i.i.d.  $X_1, \dots, X_n \sim N(0, \Sigma_{\theta_\omega})$ ,  $\omega \in B$  with

$$\delta^2 := \max_{1 \leq k \leq d} \max_{\omega \in B} \left\| \hat{T}_k(X_1, \dots, X_n) - f_k(\Sigma_{\theta_\omega}) \right\|_{L_2(\mathbb{P}_{\Sigma_{\theta_\omega}})},$$

then

$$\max_{\omega \in B} \mathbb{E}_{\Sigma_{\theta_\omega}} \frac{1}{d} \sum_{k=1}^d (\hat{T}_k - f_k(\Sigma_{\theta_\omega}))^2 \leq \delta^2.$$

- Based on  $\hat{T}_k$ , it is not hard to construct estimator  $\hat{\theta}$  such that

$$\varepsilon^2 \lesssim \max_{\omega \in B} \mathbb{E}_{\Sigma_{\theta_\omega}} \|\hat{\theta} - \theta_\omega\|^2 \lesssim \varepsilon^{2(s-1)} \delta^2,$$

implying that  $\delta \gtrsim \varepsilon^s \asymp \left(\sqrt{\frac{d}{n}}\right)^s \asymp \left(\sqrt{\frac{r}{n}}\right)^s$ .

# Estimation method: linear aggregation of plug-in estimators with different sample sizes

- $f : L(E^*, E) \mapsto \mathbb{R}, f \in C^s, s = k + \rho, k \geq 2, \rho \in (0, 1]$
- $1 \leq n_1, \dots, n_k \leq n$  sample sizes
- 

$$T_f(X_1, \dots, X_n) := \sum_{j=1}^k C_j f(\hat{\Sigma}_{n_j})$$

- The goal is to choose  $n_j$  so that the biases of plug-in estimators  $f(\hat{\Sigma}_{n_j}), j = 1, \dots, k$  “almost cancel out.”

# Bias reduction via linear aggregation of plug-in estimators (Jiao and Han (2021), for binomial model)

- $Y \sim P$  a r.v. in a Banach space  $F$  with unknown mean  $\mathbb{E}Y$
- In our case,  $F := L(E^*, E)$ ,  $Y := X \otimes X$ ,  $\mathbb{E}Y = \Sigma$
- $f : F \mapsto \mathbb{R}$
- **Goal:** estimate  $f(\mathbb{E}Y)$  based on i.i.d. observations  $Y_1, \dots, Y_n$  of  $Y$
- Let  $k \geq 2$  and let  $n/c \leq n_1 < \dots < n_k \leq n$  for some  $c > 1$ .
- Let

$$\hat{T}_f(Y_1, \dots, Y_n) := \sum_{j=1}^k C_j f(\bar{Y}_{n_j}),$$

where

$$C_j := \prod_{i \neq j} \frac{n_j}{n_j - n_i}, j = 1, \dots, k.$$



# Bias reduction via linear aggregation of plug-in estimators (Jiao and Han (2021))

## Proposition

The following properties hold:

- $\sum_{j=1}^k C_j = 1$
- $\sum_{j=1}^k \frac{C_j}{n_j^l} = 0, l = 1, \dots, k.$

## Assumption

Suppose that  $\sum_{j=1}^k |C_j| \lesssim_k 1.$

Clearly, for this assumption to hold, one needs  $n_{j+1} - n_j \asymp n.$

## Proposition

Let  $k \geq 2$  and let  $f$  be  $k$  times Fréchet differentiable with  $f^{(k)} \in \text{Lip}_\rho(F)$  for some  $\rho \in (0, 1]$ . Then

$$\left| \mathbb{E} \hat{T}_f(Y_1, \dots, Y_n) - f(\mathbb{E} Y) \right| \lesssim_{k, \rho} \|f^{(k)}\|_{\text{Lip}_\rho} \max_{1 \leq j \leq k} \mathbb{E} \|\bar{Y}_{n_j} - \mathbb{E} Y\|^{k+\rho}.$$

Moreover, if  $f$  is a polynomial of degree  $k$ , then  $\hat{T}_f(Y_1, \dots, Y_n)$  is an unbiased estimator of  $f(\mathbb{E} Y)$ .

## Proposition

Let  $k \geq 2$ ,  $\rho \in (0, 1]$  and suppose that  $f$  is  $k$  times Fréchet differentiable with  $f^{(k)} \in \text{Lip}_\rho(F)$ . Suppose also that  $\mathbb{E}\|Y\|^{k+\rho} < \infty$ . Then

$$\mathbb{E}f(\bar{Y}_n) - f(\mathbb{E}Y) = \sum_{l=1}^k \frac{\beta_{l,k}(P)}{n^l} + R,$$

where  $\beta_{l,k}(P)$ ,  $l = 1, \dots, k$  do not depend on  $n$  and

$$|R| \lesssim \|f^{(k)}\|_{\text{Lip}_\rho} \mathbb{E}\|\bar{Y}_n - \mathbb{E}Y\|^{k+\rho}.$$

If  $f$  is a polynomial of degree  $k$ , then  $R = 0$ .

# Jackknife type estimators

- $\mathcal{F}_{\text{sym}}$  the  $\sigma$ -algebra generated by symmetric functions of  $Y_1, \dots, Y_n$
- 

$$\begin{aligned}\check{T}_f(Y_1, \dots, Y_n) &:= \mathbb{E}(\hat{T}_f(Y_1, \dots, Y_n) | \mathcal{F}_{\text{sym}}) \\ &= \sum_{j=1}^k C_j U_n f(\bar{Y}_{n_j}),\end{aligned}$$

where, for  $h(Y_1, \dots, Y_m)$ ,  $m \leq n$ ,

$$\begin{aligned}(U_n h)(Y_1, \dots, Y_n) &:= \mathbb{E}(h(Y_1, \dots, Y_m) | \mathcal{F}_{\text{sym}}) \\ &= \frac{1}{\binom{n}{m}} \sum_{1 \leq j_1 < \dots < j_m \leq n} h(Y_{j_1}, \dots, Y_{j_m}).\end{aligned}$$

Note that

- $\mathbb{E} \check{T}_f(Y_1, \dots, Y_n) = \mathbb{E} \hat{T}_f(Y_1, \dots, Y_n)$
- for all  $p \geq 1$ ,

$$\begin{aligned} & \| \check{T}_f(Y_1, \dots, Y_n) - \mathbb{E} \check{T}_f(Y_1, \dots, Y_n) \|_{L_p} \\ & \leq \| \hat{T}_f(Y_1, \dots, Y_n) - \mathbb{E} \hat{T}_f(Y_1, \dots, Y_n) \|_{L_p} \\ & \lesssim_k \max_{1 \leq j \leq k} \| f(\bar{Y}_{n_j}) - \mathbb{E} f(\bar{Y}_{n_j}) \|_{L_p}. \end{aligned}$$

# Jackknife type estimators

Moreover, let

$$S_f(x, h) := f(x + h) - f(x) - \langle h, f'(x) \rangle, x, h \in F$$

be the remainder of the first order Taylor expansion.

## Proposition

For all  $p \geq 1$ ,

$$\begin{aligned} & \left\| \check{T}_f(Y_1, \dots, Y_n) - \mathbb{E} \check{T}_f(Y_1, \dots, Y_n) - \langle \bar{Y}_n - \mathbb{E} Y, f'(\mathbb{E} Y) \rangle \right\|_{L_p} \\ & \lesssim_k \max_{1 \leq j \leq k} \left\| S_f(\mathbb{E} Y, \bar{Y}_{n_j} - \mathbb{E} Y) - \mathbb{E} S_f(\mathbb{E} Y, \bar{Y}_{n_j} - \mathbb{E} Y) \right\|_{L_p}. \end{aligned}$$

# Reduction to concentration bounds

Thus, the problem reduces to:

- Bounds on

$$\mathbb{E}\|\bar{Y}_n - \mathbb{E}Y\|^{k+\rho};$$

- Bounds on

$$\|f(\bar{Y}_n) - \mathbb{E}f(\bar{Y}_n)\|_{L_p};$$

- Bounds on

$$\left\| \mathcal{S}_f(\mathbb{E}Y, \bar{Y}_n - \mathbb{E}Y) - \mathbb{E}\mathcal{S}_f(\mathbb{E}Y, \bar{Y}_n - \mathbb{E}Y) \right\|_{L_p}.$$

# Concentration bounds for sample covariance

- For  $s \geq 1$ ,

$$\mathbb{E} \|\hat{\Sigma}_n - \Sigma\|^s \lesssim_s \|\Sigma\|^s \left( \sqrt{\frac{\mathbf{r}(\Sigma)}{n}} \vee \frac{\mathbf{r}(\Sigma)}{n} \right)^s.$$

- For a Lipschitz functional  $f : L(E^*, E) \mapsto \mathbb{R}$  and for all  $\rho \geq 1$ ,

$$\|f(\hat{\Sigma}_n) - \mathbb{E}f(\hat{\Sigma}_n)\|_{L_\rho} \lesssim \|f\|_{\text{Lip}} \|\Sigma\| \left( \left( \sqrt{\frac{\mathbf{r}(\Sigma)}{n}} \vee 1 \right) \sqrt{\frac{\rho}{n}} + \frac{\rho}{n} \right).$$

- Let  $f' \in \text{Lip}_\rho(L(E^*, E))$  for some  $\rho \in (0, 1]$ . Suppose  $\mathbf{r}(\Sigma) \lesssim n$ . Then, for all  $\rho \geq 1$ ,

$$\begin{aligned} & \left\| \mathcal{S}_f(\Sigma, \hat{\Sigma}_n - \Sigma) - \mathbb{E} \mathcal{S}_f(\Sigma, \hat{\Sigma}_n - \Sigma) \right\|_{L_\rho} \\ & \lesssim \|f'\|_{\text{Lip}_\rho} \|\Sigma\|^{1+\rho} \left( \sqrt{\frac{\rho}{n}} \left( \sqrt{\frac{\mathbf{r}(\Sigma)}{n}} \right)^\rho + \left( \frac{\rho}{n} \right)^{(1+\rho)/2} + \left( \frac{\rho}{n} \right)^{1+\rho} \right). \end{aligned}$$



- Linear aggregation

$$\begin{aligned} T_f^{(1)}(X_1, \dots, X_n) \\ := \hat{T}_f(X_1 \otimes X_1, \dots, X_n \otimes X_n) &= \sum_{j=1}^k C_j f(\hat{\Sigma}_{n_j}) \end{aligned}$$

- Jackknife

$$\begin{aligned} T_f^{(2)}(X_1, \dots, X_n) \\ := \check{T}_f(X_1 \otimes X_1, \dots, X_n \otimes X_n) &= \sum_{j=1}^k C_j U_n f(\hat{\Sigma}_{n_j}) \end{aligned}$$

# A local risk bound (K (2022))

## Theorem

Suppose  $\mathbf{r}(\Sigma) \lesssim n$ . Let  $f : L(E^*, E) \mapsto \mathbb{R}$  be Lipschitz and  $k$  times Fréchet differentiable in an open ball  $U = B(\Sigma, \delta)$  of radius  $\delta > 0$  with  $\|f^{(k)}\|_{\text{Lip}_\rho(U)} < \infty$  for some  $k \geq 2$  and  $\rho \in (0, 1]$ . Suppose also that, for a sufficiently large  $C > 0$ ,  $C\|\Sigma\|\sqrt{\frac{\mathbf{r}(\Sigma)}{n}} < \delta$ . Then, for  $i = 1, 2$  and for all  $\psi \preceq \psi_1$

$$\begin{aligned} & \left\| T_f^{(i)}(X_1, \dots, X_n) - f(\Sigma) \right\|_\psi \\ & \lesssim_{k, \rho} \|f\|_{\text{Lip}} \frac{\|\Sigma\|}{\sqrt{n}} + \|f^{(k)}\|_{\text{Lip}_\rho(U)} \|\Sigma\|^{k+\rho} \left( \sqrt{\frac{\mathbf{r}(\Sigma)}{n}} \right)^{k+\rho} \\ & + \max_{2 \leq j \leq k} \|f^{(j)}(\Sigma)\| \left( \frac{\|\Sigma\|}{\sqrt{n}} \right)^j \exp \left\{ -cn \left( \frac{\delta^2}{\|\Sigma\|^2} \wedge \frac{\delta}{\|\Sigma\|} \right) \right\} \end{aligned}$$

with some constant  $c > 0$ .

## Theorem

Let  $f : L(E^*, E) \mapsto \mathbb{R}$  be Lipschitz and, for some  $k \geq 2$ , let it be  $k$  times Fréchet differentiable with  $\|f^{(k)}\|_{\text{Lip}_\rho} < \infty$  for some  $\rho \in (0, 1]$ . Then, for  $i = 1, 2$  and for all  $\psi \preceq \psi_1$ ,

$$\begin{aligned} & \left\| T_f^{(i)}(X_1, \dots, X_n) - f(\Sigma) \right\|_\psi \\ & \lesssim_S \|f\|_{\text{Lip}} \left( \sqrt{\frac{\mathbf{r}(\Sigma)}{n}} \vee 1 \right) \frac{\|\Sigma\|}{\sqrt{n}} + \|f^{(k)}\|_{\text{Lip}_\rho} \|\Sigma\|^{k+\rho} \left( \sqrt{\frac{\mathbf{r}(\Sigma)}{n}} \vee \frac{\mathbf{r}(\Sigma)}{n} \right)^{k+\rho}. \end{aligned}$$

## Theorem

Suppose that  $\mathbf{r}(\Sigma) \lesssim n$ . If  $f : L(E^*, E) \mapsto \mathbb{R}$  is  $k$  times Fréchet differentiable for some  $k \geq 2$  with  $\|f'\|_{C^1} < \infty$  and with  $\|f^{(k)}\|_{\text{Lip}_\rho} < \infty$  for some  $\rho \in (0, 1)$ , then, for all  $\beta \in [1/2, 1)$ ,

$$\begin{aligned} & \left\| T_f^{(2)}(X_1, \dots, X_n) - f(\Sigma) - \langle \hat{\Sigma}_n - \Sigma, f'(\Sigma) \rangle \right\|_{\psi_\beta} \\ & \lesssim_{k,\rho} \|f'\|_{C^1} \frac{\|\Sigma\|^{1/\beta}}{\sqrt{n}} \left( \sqrt{\frac{\mathbf{r}(\Sigma)}{n}} \right)^{1/\beta-1} + \|f^{(k)}\|_{\text{Lip}_\rho} \|\Sigma\|^{k+\rho} \left( \sqrt{\frac{\mathbf{r}(\Sigma)}{n}} \right)^{k+\rho}. \end{aligned}$$

# Normal approximation and efficiency

Let  $\sigma_f^2(\Sigma) := \text{Var}(\langle X \otimes X, f'(\Sigma) \rangle)$ . It follows that

$$\begin{aligned} & \left| \sqrt{n} \left\| T_f^{(2)}(X_1, \dots, X_n) - f(\Sigma) \right\|_{L_2} - \sigma_f(\Sigma) \right| \\ & \lesssim_{k,\rho} \|f'\|_{\mathcal{C}^1} \|\Sigma\|^2 \sqrt{\frac{\mathbf{r}(\Sigma)}{n}} + \|f^{(k)}\|_{\text{Lip}_\rho} \|\Sigma\|^{k+\rho} \sqrt{n} \left( \sqrt{\frac{\mathbf{r}(\Sigma)}{n}} \right)^{k+\rho}. \end{aligned}$$

$$\begin{aligned} & W_2 \left( \frac{\sqrt{n}(T_f^{(2)}(X_1, \dots, X_n) - f(\Sigma))}{\sigma_f(\Sigma)}, N(0, 1) \right) \lesssim_{k,\rho} \frac{\|\Sigma\|^2 \|f'(\Sigma)\|^2}{\sigma_f^2(\Sigma)} \frac{1}{\sqrt{n}} \\ & + \frac{\|f'\|_{\mathcal{C}^1} \|\Sigma\|^2}{\sigma_f(\Sigma)} \sqrt{\frac{\mathbf{r}(\Sigma)}{n}} + \frac{\|f^{(k)}\|_{\text{Lip}_\rho} \|\Sigma\|^{k+\rho}}{\sigma_f(\Sigma)} \sqrt{n} \left( \sqrt{\frac{\mathbf{r}(\Sigma)}{n}} \right)^{k+\rho}. \end{aligned}$$

# Example: an application to spectral projections

- $\Sigma_0 : \mathbb{H} \mapsto \mathbb{H}$  a covariance operator with eigenvalues  $\|\Sigma_0\| = \lambda_1 = \dots = \lambda_l > \lambda_{l+1} \geq \dots$
- $g_l := \lambda_l - \lambda_{l+1}$  the spectral gap
- $U = B(\Sigma_0; \delta) := \{A : \|A - \Sigma_0\| < \delta\}$ ,  $\delta < g_l/8$ .
- $P(A)$  the orthogonal projection onto the linear span of eigenvectors corresponding to the top  $l$  eigenvalues of  $A$
- $f(A) := \langle P(A), B \rangle$ ,  $A \in U$ ,  $\|B\|_1 < \infty$
- Then  $f \in C^\infty(U)$

# Example: an application to spectral projections

## Corollary

Let  $\gamma := \frac{\|\Sigma_0\|}{g_i}$  and suppose that  $C\gamma\sqrt{\frac{r}{n}} \leq 1$  for some  $r \geq 1$  and some constant  $C > 0$ . Then, for  $i = 1, 2$ , for all  $k \geq 1$  and for all  $\psi \preceq \psi_1$ ,

$$\sup_{\|\Sigma - \Sigma_0\| < \delta, \mathbf{r}(\Sigma) \leq r} \left\| T_{f,k}^{(i)}(X_1, \dots, X_n) - f(\Sigma) \right\|_{\psi} \lesssim_k \|B\|_1 \left( \frac{\gamma}{\sqrt{n}} + \left( \gamma \sqrt{\frac{r}{n}} \right)^{k+1} \right).$$

[K. & Lounici \(2016\)](#), [K, Löffler and Nickl \(2019\)](#): efficient estimators of linear functionals of principal components when  $\mathbf{r}(\Sigma) = o(n)$  and the top eigenvalue of  $\Sigma$  is *simple* (i.e.,  $P(\Sigma)$  is a one-dimensional spectral projection).